

Inference for partially observed Riemannian Ornstein–Uhlenbeck diffusions of covariance matrices

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Abstract

We construct a generalization of the Ornstein–Uhlenbeck processes on the cone of covariance matrices endowed with the Log-Euclidean and the Affine-Invariant metrics. Our development exploits the Riemannian geometric structure of symmetric positive definite matrices viewed as a differential manifold. We then provide Bayesian inference for discretely observed diffusion processes of covariance matrices based on an MCMC algorithm built with the help of a novel diffusion bridge sampler accounting for the geometric structure. Our proposed algorithm is illustrated with a real data financial application.

Keywords: Affine-Invariant metric; Log-Euclidean metric; Ornstein–Uhlenbeck process; Riemannian manifold; Partially observed diffusion.

1. INTRODUCTION

We are interested in Bayesian inference for diffusion processes of covariance matrices when only a discrete set of observations is available over some finite time period. Our motivation stems from financial applications where diffusions have been often adopted to describe continuous-time processes, see for example, Kalogeropoulos et al. [2010], Stramer et al. [2010], but an inferential framework to model realized covariances of asset log-returns is not available. This task is challenging not only because the marginal likelihood of the data obtained in partially observed diffusions is generally intractable, but also because dealing with covariance matrices requires models that preserve their positive definiteness. The resulting complexity of Bayesian inference via MCMC sampling algorithms necessitates the development of dynamics in $\mathcal{SP}(n)$, the curved space of symmetric positive definite matrices in $\mathbb{R}^{n \times n}$, together with sampling algorithms for diffusion bridges in $\mathcal{SP}(n)$.

There is considerable work on stochastic differential equations (SDE's) defined on positive semidefinite matrices based on Wishart processes introduced by Bru [1991] as a matrix generalisation of squared Bessel processes; see for example, Gouriéroux [2006], Gouriéroux et al. [2009], Gouriéroux and Sufana [2010], Buraschi et al. [2010], Barndorff-Nielsen and Stelzer [2007]. While Wishart processes seem natural candidates for Bayesian inference on $\mathcal{SP}(n)$, they lack geometric structure which, as will become evident in our model development, is a highly desirable property. As an example, Pfaffel [2012] notes that simulation of Wishart processes via the Euler–Maruyama method fails to always generate points on $\mathcal{SP}(n)$ so a time-adjustment algorithm is necessary.

We construct a generalisation of Ornstein–Uhlenbeck (OU) processes on $\mathcal{SP}(n)$ by noting that their dynamics are naturally specified by the Riemannian geometric structure of $\mathcal{SP}(n)$ viewed as a differential manifold endowed with the Euclidean, the Log-Euclidean (LE) and the Affine-Invariant

(AI) metrics. We emphasize the need to adopt the LE and the AI metrics which, unlike the Euclidean metric, achieves efficient sampling even from points lying near the boundary of $\mathcal{SP}(n)$. We then propose an MCMC algorithm that operates on $\mathcal{SP}(n)$ and alternates between imputing diffusivity-independent Brownian motions driving diffusion bridges between consecutive observations and approximating the likelihood with the Euler–Maruyama method adapted to the Riemannian setting via the exponential map. In particular, the construction for the AI metric required the development of a novel bridge sampling algorithm. We demonstrate our methodology with simulated and real financial data.

The essence of our construction is based on the following key points. We adopt an intrinsic point of view [Elworthy, 1982] of $\mathcal{SP}(n)$ equipped with either the LE or the AI metric. We construct a Riemannian Brownian motion as the limit of a random walk along geodesic segments [Gangolli, 1964] using the exponential map for both metrics. We then proceed to construct a $d = n(n + 1)/2$ -dimensional OU process on $\mathcal{SP}(n)$ by adding a mean reverting drift and prove that its solution is equivalent to a solution of an SDE in \mathbb{R}^d in the case of the LE metric. For the AI metric we establish existence and non-explosion of the newly proposed process. Finally, armed with the mathematical constructions, we proceed to the Bayesian estimation through a Bayesian data augmentation strategy in which the key required ingredient is the ability to sample from a diffusion bridge on $\mathcal{SP}(n)$; see Roberts and Stramer [2001].

Sampling from Brownian bridges has played an important role in Bayesian inference for diffusions. When the transition function is analytically unavailable, MCMC data augmentation sampling strategies that impute partial trajectories via bridge samplers have been used to numerically approximate the transition functions, see Elerian et al. [2001], Eraker [2001], Roberts and Stramer [2001]. The use of bridge sampling has a long history in the inference for diffusions starting from Pedersen [1995]. Recent advances include the modified diffusion bridge by Durham and Gallant [2002] and its modifications, see Golightly and Wilkinson [2008], Stramer et al. [2010], Lindström [2012] and ideas based on sequential Monte Carlo [Del Moral and Murray, 2015, Lin et al., 2010]. There has been a line of research based on ideas of Delyon and Hu [2006] that uses guided and residual proposal densities, see van der Meulen et al. [2017], Schauer et al. [2017], Whitaker et al. [2017]. Finally, a recent promising approach is based on Bladt and Sørensen [2005], see Mider et al. [2019]. We contribute to this literature by proposing a sampling strategy to sample from a diffusion bridge on $\mathcal{SP}(n)$ with AI metric which can be viewed as a guided proposal density for our MCMC sampling according to the ideas in Delyon and Hu [2006].

We investigate with both simulated and real data the performance of our proposed diffusion processes with the three metrics. We demonstrate that LE and AI metrics should be preferred to the Euclidean metric and we illustrate that both LE and AI metrics, unlike the Euclidean metric which neglects the geometric structure of $\mathcal{SP}(n)$, do not have the problem of the swelling effect [Arsigny et al., 2007] or the difficulties when sampling near the boundary of $\mathcal{SP}(n)$. We have also found that the diffusion based on the LE metric, compared with the AI metric, leads to greater anisotropy which is more evident when conditioning on matrices with eigenvalues close to zero. Our financial data example is chosen to illustrate this exact point: one can use diffusions on $\mathcal{SP}(n)$ with LE and AI metrics for pricing or portfolio construction even when the dynamics on $\mathcal{SP}(n)$ operate near the boundary.

2. RIEMANNIAN GEOMETRY FOR COVARIANCE MATRICES

2.1. Preliminary of Riemannian geometry

Smooth manifolds are motivated by the desire to extend the differentiation property to curved spaces that are more general and complicated than \mathbb{R}^d . This is achieved by considering coordinate charts, i.e. functions that map small patches of the given manifold \mathcal{M} to open sets in Euclidean space. It is then possible to define smooth curves $\gamma : [0, T] \rightarrow \mathcal{M}$ which pass through some point $\gamma(0) = P \in \mathcal{M}$, and whose velocity vectors $\dot{\gamma}(0)$ at P are known as tangent vectors constituting a vector space $T_P\mathcal{M}$, the tangent space at P . A Riemannian metric tensor g assigns to each point P on

\mathcal{M} a bilinear function g_P on $T_P\mathcal{M} \times T_P\mathcal{M}$ which is symmetric and positive definite. Smooth manifolds equipped with Riemannian metric tensors are called Riemannian manifolds and are characterised by their corresponding Riemannian metrics.

As differentiability is so special with a smooth manifold, one initially considers first order derivatives: Firstly, these include vector fields X which assign to each point $P \in \mathcal{M}$ a tangent vector $v \in T_P\mathcal{M}$ and give rise to the tangent bundle $T\mathcal{M}$, the disjoint union of all points' tangent spaces. The set of all smooth vector fields is denoted $\Gamma(T\mathcal{M})$. Secondly, differentials of smooth maps from one manifold to another which give rise to linear maps from one tangent space to another are also first order derivatives. Then, consideration moves on to second order derivatives such as the derivative of a vector field with respect to another vector field. Suppose $x = \{x^i\}_{i=1}^d$ is a local chart on an open neighbourhood U of some point P on the manifold \mathcal{M} of dimension d , the vector fields $X_i = \partial/\partial x^i$ span the tangent space $T_P\mathcal{M}$ at each $P \in U$. The covariant derivative, denoted as $\nabla_X Y$, explores how a vector field Y varies along another vector field X and the Christoffel symbols Γ_{ij}^k are functions on U defined uniquely by the relation $\nabla_{X_i} X_j = \sum_{k=1}^d \Gamma_{ij}^k X_k$, see Boothby [1986]. Moreover, using an orthonormal basis $\{E_i(P)\}_{i=1}^d$ with respect to the metric tensor g , one can simply compute the Riemannian gradient of any smooth function, i.e. $f \in C^\infty(M)$ as $\nabla_P f = \sum_{i=1}^n (E_i f)_P E_i(P)$. Here $(E_i f)_P$ can be understood as the differential of f at P in the direction of $E_i(P)$.

The connection ∇ allows us to transport a tangent vector from one tangent space to another on \mathcal{M} in a parallel manner. A vector field V along the curve γ on \mathcal{M} is said to be parallel along the curve if $\nabla_{\dot{\gamma}(t)} V = 0$ at every point on the curve [Lovett, 2010]. Furthermore, any curve $\gamma(t)$ on \mathcal{M} that satisfies $\nabla_{\dot{\gamma}(t)} \dot{\gamma}(t) = 0$ at all points on the curve is called a geodesic. They are locally defined as minimum length curves over all possible smooth curves that connect two given points on the Riemannian manifold [Caseiro et al., 2012]. The exponential map, $\text{Exp}_P : T_P\mathcal{M} \rightarrow \mathcal{M}$ computes the point at which a geodesic starting from P in the direction $\nu \in T_P\mathcal{M}$ ends after one time unit. In general, Exp_P is bijective only from a small neighbourhood $\mathcal{V} \subset T_P\mathcal{M}$ to a neighbourhood $\mathcal{U} \subset M$ of P on which the inverse map of Exp_P can be defined uniquely: this is called the logarithm map $\text{Log}_P = \text{Exp}_P^{-1}$.

We focus on the space of $n \times n$ symmetric positive definite matrices $\mathcal{SP}(n)$ which is a sub-manifold of the space of symmetric matrices $\mathcal{S}(n)$. Any metric on the space of $n \times n$ invertible matrices $\mathcal{GL}(n)$ induces a metric on $\mathcal{SP}(n)$. For example, the Frobenius inner product induces the so-called Euclidean metric g^E which, by noting that the tangent space at any point on $\mathcal{SP}(n)$ is simply $\mathcal{S}(n)$, is given by

$$g^E(S_1, S_2) = \langle S_1, S_2 \rangle_F = \text{tr}(S_1^T S_2) \text{ for } S_1, S_2 \in \mathcal{S}(n), \quad (1)$$

where tr stands for the trace operator on $\mathcal{GL}(n)$.

Since the symmetry property is not preserved under the usual matrix multiplication, Arsigny et al. [2007] proposed the use of the matrix exponential/logarithm functions:

$$P \odot Q = \exp(\log P + \log Q), \quad \lambda * P = \exp(\lambda \log P) \text{ for } P, Q \in \mathcal{SP}(n) \text{ and } \lambda \in \mathbb{R}.$$

Equipping $\mathcal{SP}(n)$ with \odot , $\mathcal{SP}(n)$ becomes an Abelian group as matrix addition is commutative. Since both matrix exponential and logarithm are diffeomorphisms, $(\mathcal{SP}(n), \odot)$ is in fact a Lie group. Moreover, we can get a vector space structure with $(\mathcal{SP}(n), \odot, *)$ since $(\mathcal{SP}(n), \odot)$ is isomorphic and diffeomorphic to $(\mathcal{S}(n), +)$. Therefore, even though $\mathcal{SP}(n)$ is not a vector space, we can identify $\mathcal{SP}(n)$ with a vector space by considering its image under the matrix logarithm. To obtain a metric, the Frobenius inner product on the Lie algebra (i.e. $T_{I_n}\mathcal{SP}(n) = \mathcal{S}(n)$) can be extended by left-translation and becomes a bi-invariant metric g^{LE} on $\mathcal{SP}(n)$. This metric is called the Log-Euclidean (LE) metric,

$$g_P^{\text{LE}}(S_1, S_2) = \langle d_P \log S_1, d_P \log S_2 \rangle_F \text{ for } S_1, S_2 \in \mathcal{S}(n), \quad (2)$$

where $d_P \log S$ is the derivative of the matrix logarithm function at $P \in \mathcal{SP}(n)$ in the direction $S \in \mathcal{S}(n)$. As the name suggests, the metric is simply the Euclidean metric in the logarithmic domain. Equipping $\mathcal{SP}(n)$ with g^{LE} , we gain invariance with respect to inversion, $g_P^{\text{LE}}(A, B) = g_P^{\text{LE}}(A^{-1}, B^{-1})$;

and invariance under similarity transform $\hat{A} = R^{-1}AR$ (where $R \in \mathcal{GL}(n)$): $g_P^{\text{LE}}(A, B) = g_P^{\text{LE}}(\hat{A}, \hat{B})$. Finally, matrices having non-positive eigenvalues are infinitely far away from any covariance matrix.

Besides the LE metric, another metric on $\mathcal{SP}(n)$, namely the Affine-Invariant (AI) metric, has been studied intensively [Pennec, 2006, Moakher, 2005, Caseiro et al., 2012]. There are many ways of defining this metric whose name arises from the group action (\star) on $\mathcal{GL}(n)$ that gives rise to Riemannian metrics invariant under this action, where

$$R \star S = RSR^T \text{ for } S \in \mathcal{S}(n), R \in \mathcal{GL}(n).$$

The AI metric g^{AI} is thus defined to satisfy

$$g_P(S_1, S_2) = g_{R \star P}(R \star S_1, R \star S_2) \text{ for } S_1, S_2 \in \mathcal{S}(n), P \in \mathcal{SP}(n) \text{ and } R \in \mathcal{GL}(n). \quad (3)$$

Choosing g_P^{AI} to be the Frobenius inner product $\langle \cdot, \cdot \rangle_F$ then defines the metric on $\mathcal{SP}(n)$:

$$g_P^{\text{AI}}(S_1, S_2) = \left\langle P^{-1/2}S_1P^{-1/2}, P^{-1/2}S_2P^{-1/2} \right\rangle_F. \quad (4)$$

Alternatively, the AI metric can be obtained from the theory of the multivariate normal distribution through the Fisher information [Skovgaard, 1984, Moakher and Zéraï, 2011].

A metric tensor can also be expressed in the form of a matrix function $G \in \mathcal{SP}(d)$ with respect to some basis. For example, the AI metric can be expressed explicitly in matrix form $G(P)$ at any $P \in \mathcal{SP}(n)$ with respect to the standard symmetric basis \mathfrak{B}_d on $\mathcal{S}(n)$, defined in equation (6):

$$G(P) = D_n^T \cdot (P^{-1} \otimes P^{-1}) \cdot D_n \text{ and } G^{-1}(P) = D_n^\dagger \cdot (P \otimes P) \cdot (D_n^\dagger)^T, \quad (5)$$

where $D_n \in \mathbb{R}^{n^2 \times n(n+1)/2}$ is a constant matrix (referred to as the duplication matrix), that satisfies $\text{vec}(P) = D_n \nu(P)$ with $\nu(P)$ containing all independent entries of P and D_n^\dagger is the Moore-Penrose inverse of D_n [Moakher and Zéraï, 2011]. Similarly to the LE metric, the AI metric is inversion-invariant and any covariance matrix is at infinite distance to any non-positive definite matrix. While the AI metric attains full affine-invariance, i.e. (3) holds for any invertible matrices, the LE metric only achieves similarity invariance, i.e. (3) only holds for orthogonal matrices.

Finally, we summarize some results about the Euclidean metric in equation (1), the Log-Euclidean metric in equation (2) and the Affine-Invariant metric in equation (4) into Table 1, which includes explicit formulae of the exponential/logarithm maps, geodesics and distance square [Arsigny et al., 2007, Pennec, 2006].

	Euclidean	Log-Euclidean	Affine-Invariant
$\text{Exp}_P(S)$	$S + P$	$\exp(\log P + d_P \log S)$	$P^{1/2} \star \exp(P^{-1/2} \star S)$
$\text{Log}_P(Q)$	$Q - P$	$d_{\log P} \exp(\log Q - \log P)$	$P^{1/2} \star \log(P^{-1/2} \star Q)$
$\gamma_{(P,Q)}(t)$	$P + t(Q - P)$	$\exp(\log P + t(\log Q - \log P))$	$P^{1/2} \star \exp(P^{-1/2} \star tQ)$
$d^2(P, Q)$	$\ Q - P\ _F^2$	$\ \log Q - \log P\ _F^2$	$\ \log(P^{-1/2} \star Q)\ _F^2$

Table 1: Explicit formulae of exponential map, logarithm map, geodesic and distance square for the Euclidean, Log-Euclidean and Affine-Invariant metrics. For the Log-Euclidean and Affine-Invariant cases, these will be denoted as Exp^{LE} , Exp^{AI} , Log^{LE} , Log^{AI} and d_{LE} , d_{AI} , respectively.

Let us fix an orthonormal basis $\mathfrak{B}_d = \{S_i\}_{i=1}^d$ with respect to the Frobenius inner product on the tangent space $\mathcal{S}(n)$ of $\mathcal{SP}(n)$, where $d = n(n+1)/2$:

$$S_i = e_{ii}^{(n)} \quad \text{for } 1 \leq i \leq n \quad (6)$$

$$S_{n+1} = (e_{21}^{(n)} + e_{12}^{(n)})/\sqrt{2}, S_{n+2} = (e_{31}^{(n)} + e_{13}^{(n)})/\sqrt{2}, S_{n+3} = (e_{32}^{(n)} + e_{23}^{(n)})/\sqrt{2}, \dots$$

Here, $\{S_i\}_{i=1}^n$ has all entries zero except the j th entry on the diagonal being one. The remaining $\{S_i\}_{i=n+1}^d$ are obtained by adding, with $i > j$, the single-entry matrix $e_{ij}^{(n)}$ with one at the (i, j) th entry and zero elsewhere to its transpose and dividing by $\sqrt{2}$ so that it has unit Frobenius norm. We call \mathfrak{B}_d the standard symmetric basis of $\mathcal{S}(n)$.

2.2. The importance of Riemannian geometry to $\mathcal{SP}(n)$

We discuss two major reasons that necessitate the use of Riemannian geometry: easier sampling close to the boundaries of $\mathcal{SP}(n)$ and no swelling effects. One may additionally argue that other properties, such as inversion-invariance and similarity-invariance for the LE and AI metrics, and affine-invariance for the AI metric, may be useful in calculations for complicated computational algorithms.

Although the Frobenius inner product on $\mathcal{SP}(n) \subset \mathcal{GL}(n)$ is simple, it is problematic because non-covariance matrices are only a finite distance away from covariance matrices. As Table 1 illustrates, LE and AI metrics do not suffer from this problem as the involvement of the matrix logarithm guarantees that non-covariance matrices are at infinite distance from any point on $\mathcal{SP}(n)$. Therefore, they avoid the undesirable inequality constraints that are required in the Frobenius induced geometry to ensure positive definiteness and whose number grows quadratically with n . As will become evident in our simulation experiments, this turns out to be a highly desirable property because it facilitates sampling close to the boundary of $\mathcal{SP}(n)$.

The determinant of a covariance matrix measures the dispersion of the data points from a multivariate normal distribution. For the Euclidean metric, the geodesic connecting two fixed points often contains points with a larger determinant than the two fixed points, and the difference can get extremely large whenever the fixed points lie near the boundary of $\mathcal{SP}(n)$. This problem is referred to as the swelling effect; see for example, Arsigny et al. [2007], Dryden et al. [2009], Jung et al. [2015]. In many contexts, the swelling effect is described as undesirable because the level of dispersion should remain close to the given information obtained by the observations of covariance matrices; see for example, Arsigny et al. [2007], Chéfd’Hotel et al. [2004], Fletcher and Joshi [2007], Tschumperle and Deriche [2001]. The LE and AI metrics avoid this swelling effect. Moreover, points on LE and AI geodesics at corresponding times have the same determinants and these determinants are the result of linear interpolation in the logarithmic domain; this can be proved by following similar lines as in Arsigny et al. [2007].

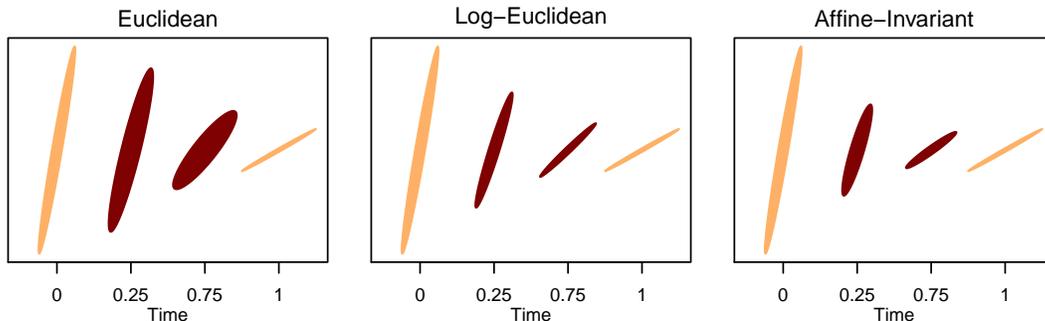


Figure 1: Comparison of three metric tensors on $\mathcal{SP}(2)$: Red ellipses represent two points at time $t = 0.25, 0.75$ on the geodesic connecting P_0 at $t = 0$ and P_1 at $t = 1$. The length of the axes are the square root of the eigenvalues.

A visual illustration of the swelling effect is provided in Figure 1 where two intermediate points on the geodesic connecting $P_0 = \begin{pmatrix} 0.4 & 0.3 \\ 0.3 & 0.4 \end{pmatrix}$ at $t = 0$ and $P_1 = \begin{pmatrix} 1 & 0.1 \\ 0.1 & 0.02 \end{pmatrix}$ at $t = 1$ are shown for each metric. Notice the swelling effect in the case of the Euclidean metric and also that the geodesic of the LE metric has points with exaggerated anisotropy; for more details on this phenomenon see Arsigny et al. [2006], Dryden et al. [2009].

3. STOCHASTIC PROCESSES ON $\mathcal{SP}(n)$

3.1. Overview

SDEs on manifolds present additional complications over the Euclidean setting which we deal with in three stages: firstly, we introduce the general notion of SDEs on manifolds. Secondly, we discuss Brownian Motion (definition, local representation, horizontal lift, construction via Euler-Maruyama with the exponential map and non-explosion) in subsection 3.2. and thirdly we proceed to the Ornstein–Uhlenbeck class in subsection 3.3..

Since the curvature of Riemannian manifolds makes direct use of Euclidean stochastic analysis prohibitively hard, a common strategy in stochastic analysis is to adopt the extrinsic view that involves embedding the manifold in a higher dimensional Euclidean space using the Nash embedding theorem; see for example, Hsu [2002], Elworthy [1982], Gangolli [1964]. The approach benefits from existing theory for SDE's on Euclidean space but a suitable coordinate system is often not explicitly available or too inconvenient, so this approach has limited practical use.

On a Riemannian manifold \mathcal{M} of dimension d , Hsu [2002] writes the SDE driven through vector fields V_1, \dots, V_l by an \mathbb{R}^l -valued semi-martingale Z_t (with initial condition $P \in \mathcal{F}_0$) as

$$dX_t = \sum_{i=1}^l V_i(X_t) \circ dZ_t^i \quad (X_0 = P). \quad (7)$$

Z_t^i could be a deterministic component, such as time, or a stochastic component, such as a Brownian motion: this corresponds to the usual distinction of some V_i as drift and other V_i as diffusivity. We call X_t an \mathcal{M} -valued semi-martingale defined up to a stopping time τ if it satisfies

$$f(X_t) = f(P) + \sum_{i=1}^l \int_0^t V_i f(X_s) \circ dZ_s^i \quad 0 \leq t < \tau, \quad f \in C^\infty(\mathcal{M}), \quad (8)$$

where the integrals above are in the Stratonovich sense, and converting them to Ito sense yields:

$$f(X_t) = f(P) + \sum_{i=1}^l \int_0^t V_i f(X_s) dZ_s^i + \frac{1}{2} \sum_{i,j=1}^l \int_0^t (\nabla_{V_j} V_i) f(X_s) d\langle Z^i, Z^j \rangle_s. \quad (9)$$

The additional terms in equation (9) arise from the non-trivial chain rule in the Ito case. The Stratonovich representation in equation (8) brings simplicity and is invariantly defined whence it is usually preferred for SDE's on manifolds [Elworthy, 1982].

3.2. Brownian motion class

Since the infinitesimal generator of Brownian motion on Euclidean space is $\Delta/2$, with Δ the usual Laplace operator, Brownian motion on a Riemannian manifold \mathcal{M} can also be defined as a diffusion process generated by $\Delta_{\mathcal{M}}/2$ where $\Delta_{\mathcal{M}}$ denotes the Laplace-Beltrami operator. The resulting Brownian Motion can be expressed in local coordinates using the standard Brownian motion B_t on \mathbb{R}^d by writing $\Delta_{\mathcal{M}}$ in local coordinates:

$$dX_t^i = \sum_{j=1}^d (G^{-1/2})_{ij}(X_t) dB_t^j - \frac{1}{2} \sum_{k,l=1}^d (G^{-1})_{lk}(X_t) \Gamma_{kl}^i(X_t) dt \quad 1 \leq i \leq d, \quad (10)$$

where $(G^{-1})_{ij}$ is the (i, j) -entry of G^{-1} which is the matrix form of the metric tensor g of the manifold \mathcal{M} [Hsu, 2002, Elworthy, 1982]. We choose to instead adopt the intrinsic viewpoint that studies Riemannian manifolds via their metric or connection which enables us to write down less cumbersome SDEs with more readily interpretable parameters. Thus, let us introduce a frame $u : \mathbb{R}^d \rightarrow T_P \mathcal{M}$

at $P \in \mathcal{M}$ (this is an isomorphism of vector spaces with inner product) and the frame bundle $\mathfrak{F}(\mathcal{M})$ which is the collection of all frames for all $P \in \mathcal{M}$. Let us fix the standard basis $\{e_i\}_{i=1}^d$ on \mathbb{R}^d and express u in local coordinates $\{E_1, \dots, E_d, e_1, \dots, e_d\}$ in some neighbourhood U of $P \in \mathcal{M}$ as $u = (P, \zeta)$ with $\zeta = (\zeta_j^i) \in \mathbb{R}^{d \times d}$ the coefficients with respect to the orthonormal basis $\{E_i(P)\}_{i=1}^d$ on $T_P\mathcal{M}$. This means that for any vector $e \in \mathbb{R}^d$ with coordinates ϵ_i , i.e. $e = \sum_{i=1}^d \epsilon_i e_i \in \mathbb{R}^d$:

$$u(e) = \sum_{i,j=1}^d \epsilon_i \zeta_j^i E_i(P) \in T_P\mathcal{M}.$$

Moreover, $\mathfrak{F}(\mathcal{M})$ is again a smooth manifold, and the canonical projection map $\pi : \mathfrak{F}(\mathcal{M}) \rightarrow \mathcal{M}$ is smooth [Hsu, 2002]. If u_t is a smooth curve on $\mathfrak{F}(\mathcal{M})$ and for each $e \in \mathbb{R}^d$ the vector field $u_t(e)$ of \mathcal{M} is parallel along the curve $\pi(u_t)$, u_t is called horizontal curve on $\mathfrak{F}(\mathcal{M})$. For any smooth curve γ on \mathcal{M} , there is a corresponding horizontal curve on $\mathfrak{F}(\mathcal{M})$ (unique up to choice of initial condition u_0) which is referred to as the horizontal lift of γ . This definition carries over to the horizontal lift of a tangent vector on \mathcal{M} . We define the anti-development of γ as

$$w_t = \int_0^t u_s^{-1}(\dot{\gamma}(s)) ds,$$

where u_t is the horizontal lift of $\gamma(t)$ on \mathcal{M} . While this anti-development w_t is guaranteed to exist and is uniquely defined up to the initial conditions u_0 and $\gamma(0)$, its computation is often difficult. To address this issue, we establish the following:

Proposition 1. *On $\mathcal{SP}(n)$, in the case of the LE metric, the horizontal lift of a smooth curve can be explicitly expressed in local coordinates. For the AI metric, a first order Euler approximation for the horizontal lift of a geodesic can be explicitly computed.*

More details about the stochastic development on $\mathcal{SP}(n)$ are discussed in Proposition S3 and Corollary S1 in the Supplementary Material.

By carrying out a similar process to the case of a smooth curve, we obtain a corresponding horizontal semi-martingale U_t on the frame bundle and an anti-development W_t on \mathbb{R}^d to a semi-martingale X_t on \mathcal{M} . Up to a choice of initial conditions, this relationship is one-to-one. The process of transforming W_t to X_t is called stochastic development, see Elworthy [1982]. One particularly important result is that one can define Riemannian Brownian motion on \mathcal{M} with the connection ∇ by having its anti-development W_t be the standard Euclidean Brownian motion. On the sphere S^2 , stochastic development is intuitively described as “rolling without slipping”: if we have a path of Brownian motion W_t on a flat paper (this paper acts as the tangent plane of S^2), rolling the sphere along the path of W_t without slipping results in a trajectory on S^2 which turns out to be a path of the Riemannian Brownian motion on S^2 .

Yet another alternative construction of Riemannian Brownian Motion uses the Euler–Maruyama approximation, which employs the exponential map:

$$X_{t+\delta t} = \text{Exp}_{X_t} \left\{ \sum_{i=1}^d (B_{t+\delta t}^i - B_t^i) E_i(X_t) \right\} \quad \text{for } \delta t > 0; \quad (11)$$

see for example, Baxendale [1976], Manton [2013], McKean et al. [1960]. As $\delta t \rightarrow 0$, X_t converges to the Riemannian Brownian motion in distribution [Gangolli, 1964]. Since $\mathcal{SP}(n)$ endowed with either the LE or the AI metric is geodesically complete (i.e. the exponential map is a global diffeomorphism), the approximation method in equation (11) becomes more convenient and efficient in our case.

The transition density function $p_{\mathcal{M}}(s, P; t, Q)$ of the Riemannian Brownian motion exists but usually no explicit expression is available, while on the Euclidean space it is simply the Gaussian distribution [Hsu, 2002, Elworthy, 1982]. On the Euclidean space, Brownian motion does not explode in finite time, and if this holds in the Riemannian setting, that is

$$\int_{\mathcal{M}} p_{\mathcal{M}}(0, P; t, Q) dQ = 1 \quad \text{for all } P \in \mathcal{M} \text{ \& } 0 < t < \infty,$$

then the manifold \mathcal{M} is said to be stochastically complete. It turns out that $\mathcal{SP}(n)$ equipped either with the LE or the AI metric is also stochastically complete, see Proposition S3 in the Supplementary Material.

3.3. Ornstein–Uhlenbeck (OU) class

Adopting the intrinsic point of view, we present a construction of an OU class of processes on $\mathcal{SP}(n)$ for both the LE and the AI metrics. In analogy with the Euclidean OU process, we start with Brownian motion and add a mean-reverting drift which pushes the process toward the point of attraction. In the Euclidean setting, this drift is simply given as the gradient of the squared distance between the process and the point of attraction. We translate this idea to manifolds by using the covariant derivative in the place of the Euclidean gradient. This is similar to the treatment of the drift term by Staneva and Younes [2017] for shape manifolds.

Let us define the OU process on $\mathcal{SP}(n)$ to be the solution of the following SDE with model parameters $\theta \in \mathbb{R}_{>0}$, $M \in \mathcal{SP}(n)$ and $\sigma \in \mathbb{R}_{>0}$:

$$dX_t = -\frac{\theta}{2} \nabla_{X_t} \{d^2(X_t, M)\} dt + F_{X_t}(\sigma dB_t) \quad (X_0 = P). \quad (12)$$

This uses the smooth function $F : \mathcal{SP}(n) \times \mathbb{R}^d \rightarrow \Gamma(T\mathcal{SP}(n))$ defined as:

$$F_Q(e) = \sum_{i=1}^d \epsilon_i E_i(Q) \quad \text{with} \quad e = \sum_{i=1}^d \epsilon_i e_i \in \mathbb{R}^d, \quad Q \in \mathcal{SP}(n) \quad \text{and} \quad d = \frac{n(n+1)}{2}. \quad (13)$$

The following Proposition demonstrates that the covariant derivative chosen for the drift in the SDE (12) is explicitly computable; the proof is presented in Proposition S2 (Supplementary Material).

Proposition 2 (Riemannian gradient of distance squared on $\mathcal{SP}(n)$).

(i) (LE metric). The set $\mathfrak{B}_d^{LE} = \{E_i^{LE}\}_{i=1}^d$ is an orthonormal frame on the tangent bundle $T\mathcal{SP}(n)$, where for any $P \in \mathcal{SP}(n)$:

$$E_i^{LE}(P) = d_{\log P} \exp S_i = (d_P \log)^{-1} S_i \quad \text{for } 1 \leq i \leq d.$$

Moreover, the Riemannian gradient of distance squared for any fixed point $Q \in \mathcal{SP}(n)$ is $\nabla_P \{d_{LE}^2(P, Q)\} = -2(d_P \log)^{-1}(\log Q - \log P) = -2 \text{Log}_P^{LE}(Q)$.

(ii) (AI metric). The set $\mathfrak{B}_d^{AI} = \{E_i^{AI}\}_{i=1}^d$ is an orthonormal frame on the tangent bundle $T\mathcal{SP}(n)$, where for any $P \in \mathcal{SP}(n)$:

$$E_i^{AI}(P) = P^{1/2} \star S_i \quad \text{for } 1 \leq i \leq d.$$

Moreover, the Riemannian gradient of distance squared for any fixed point $Q \in \mathcal{SP}(n)$ is $\nabla_P \{d_{AI}^2(P, Q)\} = -2 \sum_{i=1}^d \langle \log(P^{-1/2} Q P^{-1/2}), S_i \rangle_F E_i^{AI}(P) = -2 \text{Log}_P^{AI}(Q)$.

For simulation, we employ the Euler–Maruyama method adapted to the Riemannian setting via the exponential map, i.e. using the standard Euler–Maruyama method on the tangent space:

$$X_{t+\delta_t} = \text{Exp}_{X_t} \left\{ -\frac{\theta}{2} \nabla_{X_t} \{d^2(X_t, M)\} \delta_t + \sum_{j=1}^d (B_{t+\delta_t}^j - B_t^j) \sigma E_j(X_t) \right\} \quad \text{for } \delta_t > 0. \quad (14)$$

The selection of the basis fields plays an important role as they represent the horizontal lift of X_t locally when using the piece-wise approximation method in equation (14).

For the LE metric, we define a global isometric diffeomorphism \mathfrak{h} , that allows us to constructively identify $\mathcal{SP}(n)$ with \mathbb{R}^d . Additionally, it characterizes the OU class of processes on $\mathcal{SP}(n)$ equipped

with the LE metric as the image of the standard OU process on \mathbb{R}^d under \mathfrak{h} , see Theorem 1 in which the proof is presented in Theorem S1 (Supplementary Material). In turn, this permits establishing existence, uniqueness and non-explosion of the OU class for the LE metric.

We define $\mathfrak{h} = (\mathfrak{h}_j) : \mathcal{SP}(n) \rightarrow \mathbb{R}^d$ with $e = \sum_{j=1}^d \epsilon_j e_j \in \mathbb{R}^d$ and $P \in \mathcal{SP}(n)$ as follows:

$$\mathfrak{h}_j(P) = \langle \log P, S_j \rangle_F \quad (1 \leq j \leq d) \quad \text{and} \quad \mathfrak{h}^{-1}(e) = \exp \left(\sum_{j=1}^d \epsilon_j S_j \right). \quad (15)$$

Theorem 1. *Consider the following SDE on $\mathcal{SP}(n)$ endowed with the LE metric :*

$$dX_t = V(X_t) dt + F_{X_t}(a(X_t)) dB_t \quad (X_0 = P), \quad (16)$$

where V is a smooth vector field with smooth function $a : \mathcal{SP}(n) \rightarrow \mathbb{R}^{d \times d}$, B_t is \mathbb{R}^d -valued Brownian Motion and the function F is as defined in equation (13). Then the problem of solving the SDE (16) on $\mathcal{SP}(n)$ is the same as solving the following SDE on \mathbb{R}^d :

$$dx_t = v(x_t) dt + \tilde{a}(x_t) dB_t \quad (x_0 = p), \quad (17)$$

- (i) $p = \mathfrak{h}(P)$, $x_t = \mathfrak{h}(X_t)$ for all $t > 0$ and $\tilde{a} : \mathbb{R}^d \rightarrow \mathbb{R}^{d \times d}$, $e \mapsto a(\mathfrak{h}^{-1}(e))$,
- (ii) Function $v = (v^{(j)})$ with $v^{(j)} \in C^\infty(\mathbb{R}^d)$, $x_t \mapsto \langle (d_{X_t} \log) V(X_t), S_j \rangle_F$ ($1 \leq j \leq d$).

Solutions of (16) and (17) are in one-to-one correspondence. Therefore, the conditions for existence and uniqueness of the solution for the SDE (16) depend directly on the requirements that the drift and diffusivity of the SDE (17) satisfy on the Euclidean space, e.g. continuity and local Lipschitzness. Indeed, equating the drift and diffusivity of the SDE (16) with our OU process, the SDE (17) turns out to be a standard OU process on \mathbb{R}^d . Thus, most favourable properties that the OU process has on the Euclidean space will carry over to $\mathcal{SP}(n)$, such as existence and uniqueness of the solution and ergodicity. The transition probability density is explicitly available up to the Jacobian term involving the derivative of the matrix exponential.

Corollary 1. *SDE (12) has a unique solution and gives rise to an ergodic diffusion process on $\mathcal{SP}(n)$ in the LE case.*

We conclude this section by establishing equivalent results for the AI case in a non-constructive manner. While there is no simple diffeomorphism corresponding to \mathfrak{h} , otherwise equivalent results can be obtained for the AI case:

Proposition 3. *The existence and uniqueness theorem in Elworthy [1982, Theorem 2E, Page 121] is applicable to the OU process on $\mathcal{SP}(n)$ equipped with the AI metric. Moreover, this diffusion process is also non-explosive, see Elworthy [1982, Corollary 6.1, Page 131].*

4. BAYESIAN PARAMETER ESTIMATION

We now focus on the Bayesian estimation of the parameters of the OU diffusion processes on $\mathcal{SP}(n)$ when observations are collected at low frequency. We adopt the data augmentation MCMC computational strategy introduced by Roberts and Stramer [2001] which requires data imputation through sampling from a diffusion bridge. We need to build diffusion bridge samplers that operate on $\mathcal{SP}(n)$ which, unlike the Euclidean case, have not been studied before. A common approach used in manifolds is to use embedding or local charts followed by an appropriate Euclidean method; see for example, Ball et al. [2008], Sommer et al. [May 2017], Staneva and Younes [2017], but this strategy is unsuitable when transitioning between charts is required and charts can be cumbersome to work with. We therefore develop a diffusion bridge sampler exploiting the exponential map and adopting an intrinsic viewpoint. In fact, by using Corollary 1 we can translate any existing method

for the OU process from the Euclidean to the LE setting, so in the remainder of this section we will focus only on the AI metric where no such result exists. To deal with the data augmentation problem it is either assumed that X_t has constant diffusivity, or that X_t is transformed to a process of constant diffusivity, or existence of a process that is absolutely continuous to X_t and its corresponding transition probability density need to be derived. In the case of the AI metric, at first glance the SDE seems to have constant diffusivity. However, due to the presence of curvature, the diffusivity part does depend on the position of the process X_t , hence straightforward algorithms from literature are not applicable. In fact, by looking at the local coordinate expression of the Riemannian Brownian motion on $\mathcal{SP}(n)$ equipped with the AI metric (substituting G^{-1} in equation (5) and Christoffel symbols given in Lemma 1 into equation (10)), the dependence of the diffusivity on X_t is evident as is the complexity of the resulting expressions. Furthermore, attempting to sample from this Euclidean SDE using the standard Euler-Maruyama method will lead to symmetric but non-positive definite matrices.

On the Euclidean space, Delyon and Hu [2006] and Schauer et al. [2017] suggest adding an extra drift term which guides the SDE solution toward the correct terminal point and leaves its law absolutely continuous with respect to the law of the original conditional diffusion process; the Radon-Nikodym derivative is available explicitly. This results in an easier simulation, much better MCMC mixing rate of convergence and no difficulty of computing acceptance probability when updating the proposal bridge. The additional drift is the gradient of the logarithm of the transition density of an auxiliary process \tilde{X}_t which must have explicitly available transition probability density. In the manifold setting, we are aware of one attempt to use the above approach to sample diffusion bridges through local coordinates by Sommer et al. [May 2017], but using the exponential map in this context is new. Motivated by these ideas, we construct a methodology that allow us to sample a diffusion bridge on $\mathcal{SP}(n)$ equipped with the AI metric using a guided proposal process.

We need to choose a proposal diffusion process, which has both an explicit transition probability density and an analytically tractable gradient of the log transition probability density due to the requirement in obtaining the additional drift in the guided proposals SDE. The Wishart process is inappropriate because it does not have a closed form for its Riemannian gradient so we choose a diffusion process \tilde{X}_t on $\mathcal{SP}(n)$ with transition probability the Riemannian Gaussian distribution [Said et al., 2017] given as

$$p(s, \tilde{X}_s; t, \tilde{X}_t) = \frac{1}{K_n(\sigma)} \exp\left(-\frac{d_{\text{AI}}^2(\tilde{X}_s, \tilde{X}_t)}{2(t-s)\sigma^2}\right),$$

where $K_n(\sigma)$ is the normalising constant that depends only on σ and n . We emphasize that explicit availability of the guided proposal SDE is not actually a pre-requisite for the guided proposals algorithm. Indeed, the proposal Markov process \tilde{X}_t exists but its SDE form is not explicitly available [Baxendale, 1976].

Consider the OU process X_t on $\mathcal{SP}(n)$ equipped with the AI metric with its law \mathbb{P}_t and assume that sampling from the target diffusion bridge $X_t^* = \{X_t, 0 \leq t \leq T \mid X_0 = U, X_T = V\}$ with its corresponding law \mathbb{P}_t^* is required. We introduce the guided proposal X_t^\diamond which is the solution of the following SDE with its law \mathbb{P}_t^\diamond :

$$dX_t^\diamond = \left(\theta \text{Log}_{X_t^\diamond}^{\text{AI}} M + \frac{\text{Log}_{X_t^\diamond}^{\text{AI}} V}{T-t} \right) dt + F(X_t^\diamond)(\sigma dB_t) \quad (X_0^\diamond = U). \quad (18)$$

We then show that \mathbb{P}_t^* and \mathbb{P}_t^\diamond are equivalent up to time T with the aid of Lemma 1 in Theorem 2; proofs are shown in Lemma S1 and Theorem S2 respectively (Supplementary Material).

Lemma 1. *The Christoffel symbols at any $P \in \mathcal{SP}(n)$ with respect to the basis $\mathfrak{B}_d^{\text{AI}}$ do not depend on P and are given by $\Gamma_{ij}^k(P) = \langle -(S_i S_j + S_j S_i), S_k \rangle_F / 2$.*

Theorem 2. *For $t \in [0, T)$ the laws $\mathbb{P}_t, \mathbb{P}_t^*$ and \mathbb{P}_t^\diamond are absolutely continuous. Let $p(t, X_t; T, V)$ be the true (unknown) transition density of moving from X_t at time t to V at time T and let $X_{[0:t]}^\diamond$ be*

the path of X_t^\diamond from time 0 to t . Then

$$\frac{d\mathbb{P}_t}{d\mathbb{P}_t^\diamond}(X_{[0:t]}^\diamond) = \frac{\exp(f(X_0^\diamond))}{\exp(f(X_t^\diamond))} \exp \left\{ \Phi(t, X_{[0:t]}^\diamond) + \phi(t, X_{[0:t]}^\diamond) \right\}, \quad (19)$$

$$\frac{d\mathbb{P}_t^*}{d\mathbb{P}_t^\diamond}(X_{[0:t]}^\diamond) = \frac{p(t, X_t^\diamond; T, V) \exp(f(X_0^\diamond))}{\exp(f(X_t^\diamond)) p(0, U; T, V)} \exp \left\{ \Phi(t, X_{[0:t]}^\diamond) + \phi(t, X_{[0:t]}^\diamond) \right\}, \quad (20)$$

where f , ϕ and Φ are

$$\begin{aligned} f(X_t^\diamond) &= -\frac{d_{AI}^2(X_t^\diamond, V)}{2\sigma^2(T-t)} = -\frac{\|\log((X_t^\diamond)^{-1/2}V(X_t^\diamond)^{-1/2})\|_F^2}{2\sigma^2(T-t)}, \\ \phi(t, X_{[0:t]}^\diamond) &= \sum_{i,j=1}^d \int_0^t \frac{(\zeta_j^i(X_s^\diamond))^2}{2(T-s)} ds, \end{aligned} \quad (21)$$

$$\begin{aligned} \Phi(t, X_{[0:t]}^\diamond) &= \int_0^t \left[\theta g_{X_s^\diamond}^{AI} \left(\text{Log}_{X_s^\diamond}^{AI} M, \frac{\text{Log}_{X_s^\diamond}^{AI} V}{\sigma^2(T-s)} \right) \right. \\ &\quad \left. + \frac{1}{2} \sum_{i,j,r=1}^d g_{X_s^\diamond}^{AI} \left(\zeta_j^i(X_s^\diamond) \left(\sum_{l=1}^d \zeta_l^i(X_s^\diamond) \Gamma_{jl}^r + (E_j^{AI} \zeta_r^i)_{X_s^\diamond} \right) E_r^{AI}(X_s^\diamond), \frac{\text{Log}_{X_s^\diamond}^{AI} V}{T-s} \right) \right] ds, \end{aligned} \quad (22)$$

with Γ_{jl}^r given in Lemma 1. The functions $\zeta = (\zeta_i^i)$ are the coefficients with respect to the basis \mathfrak{B}_d^{AI} in the expression of the horizontal lift in local coordinates.

The Radon-Nikodym derivatives in equations (19)–(20) can not be computed explicitly due to the presence of ζ . However, we can approximate ϕ and Φ in Theorem 2 based on the approximation of ζ , see Corollary S1 (Supplementary Material). As a result, Remark 1 is a crucial step to make our proposed algorithm practicable; detailed calculation of the approximation of ϕ and Φ are presented in Corollary S2 (Supplementary Material).

Remark 1. Given that X_t^\diamond is simulated using the piece-wise approximation method in equation (14) with sufficiently small time step, that is we have a simulation path $\{X_{t_k}^\diamond = y_{t_k}^\diamond\}_{k=0}^{m+1}$ such that $\max\{t_{k+1} - t_k\}_{k=0}^m$ is sufficiently small and $0 = t_0 < \dots < t_{m+1} = t$, the functions ϕ in equation (21) and Φ in equation (22) can be approximated as follows:

$$\begin{aligned} \phi(t, X_{[0:t]}^\diamond) &\approx \frac{d}{2} \log \frac{T-t}{T}, \\ \Phi(t, X_{[0:t]}^\diamond) &\approx \sum_{k=0}^m \frac{t_{k+1} - t_k}{T - t_k} \left[\frac{\theta \langle \log \{ (y_{t_k}^\diamond)^{-1/2} M (y_{t_k}^\diamond)^{-1/2} \}, \log \{ (y_{t_k}^\diamond)^{-1/2} V (y_{t_k}^\diamond)^{-1/2} \} \rangle_F}{\sigma^2} \right. \\ &\quad \left. + \frac{\langle \Gamma, \log((y_{t_k}^\diamond)^{-1/2} V (y_{t_k}^\diamond)^{-1/2}) \rangle_F}{2} \right], \end{aligned}$$

with $\Gamma = \sum_{i,r=1}^d \Gamma_{ii}^r S_r$ and Christoffel symbols Γ_{ii}^r are given in Lemma 1.

We expect that taking the limit $t \uparrow T$ in equation (20) will follow along similar lines to those in Delyon and Hu [2006] and Schauer et al. [2017] so that the following holds

$$\frac{d\mathbb{P}_T^*}{d\mathbb{P}_T^\diamond}(X_{[0:T]}^\diamond) = \frac{\mathcal{H}_{n,T}}{p(0, U; T, V)} \exp \left\{ f(X_0^\diamond, \sigma^2) + \Phi(T, X_{[0:T]}^\diamond) - \frac{d}{2} \log \sigma^2 \right\},$$

where $\mathcal{H}_{n,T}$ is a fixed constant that depends only on the dimension n and the terminal time T . While we do not present the full argument, we do, however, provide a careful numerical validation in Section 5..

Suppose we have discretely observed data $\mathfrak{D} = \{X_{t_j} = y_j\}_{j=0}^N$ at observation times $t_0 = 0 < t_1 < \dots < t_N = T$, where the diffusion process X_t is the OU process X_t on $\mathcal{SP}(n)$. We aim to sample from the posterior distribution of $\Theta = \{\theta, M, \sigma^2\}$. We set $\mu = \sum_{i=1}^d \mu^{(i)} e_i = \mathfrak{h}(M)$, as defined in equation (15), and the prior distributions of $\{\theta, \mu, \sigma^2\}$ as π_0^θ , π_0^μ and π_0^σ respectively. The key step involves first imputing suitable $m_j - 1$ data points between the j th consecutive observations in a way that they are independent of the diffusivity, and then using the Riemannian adapted Euler–Maruyama approximation for the likelihood. We choose to use random walk symmetric proposal distributions with suitable choices of step size $q(\tilde{\theta}|\theta)$, $q(\tilde{\mu}|\mu)$ and $q(\tilde{\sigma}^2|\sigma^2)$ for $\{\theta, \mu, \sigma^2\}$ respectively.

Algorithm 1 (Guided proposals on $\mathcal{SP}(n)$).

1. (Iteration $k = 0$). Choose starting values for Θ and sample standard Brownian motions W_j , independently for $1 \leq j \leq N$, each covering the time interval $t_j - t_{j-1}$, and set $B_j^{(0)} = W_j$.
2. (Iteration $k \geq 1$).

(a) Update B_j independently ($1 \leq j \leq N$): sample the proposal \tilde{W}_j and obtain $\tilde{Y}_{[t_{j-1}, t_j]}$ from $\{\tilde{W}_j, \theta_{k-1}, M_{k-1}, \sigma_{k-1}^2, \mathfrak{D}\}$ and $Y_{[t_{j-1}, t_j]}$ from $\{B_j^{(k-1)}, \theta_{k-1}, M_{k-1}, \sigma_{k-1}^2, \mathfrak{D}\}$ using equation (14) to approximately solve the SDE (18); then accept \tilde{W}_j with probability $\alpha^{(B)} = \exp\left[\Phi(t_j - t_{j-1}, \tilde{Y}_{[t_{j-1}, t_j]}) - \Phi(t_j - t_{j-1}, Y_{[t_{j-1}, t_j]})\right]$

(b) Update σ^2 : sample proposal $\tilde{\sigma}^2$ from $q(\sigma^2|\sigma_{k-1}^2)$ and obtain $\tilde{Y}_{[t_{j-1}, t_j]}$ from $\{B_j^{(k)}, \theta_{k-1}, M_{k-1}, \tilde{\sigma}^2, \mathfrak{D}\}$ and $Y_{[t_{j-1}, t_j]}$ from $\{B_j^{(k)}, \theta_{k-1}, M_{k-1}, \sigma_{k-1}^2, \mathfrak{D}\}$ using equation (14) to approximately solve the SDE (18); then accept $\tilde{\sigma}^2$ with probability

$$\alpha^{(\sigma)} = \frac{\pi_0^\sigma(\tilde{\sigma}^2)}{\pi_0^\sigma(\sigma_{k-1}^2)} \frac{\prod_{j=1}^N \exp\left\{\Phi(t_j - t_{j-1}, \tilde{Y}_{[t_{j-1}, t_j]}) + f(\tilde{Y}_{t_{j-1}}, \tilde{\sigma}^2) - \frac{d}{2}\tilde{\sigma}^2\right\}}{\prod_{j=1}^N \exp\left\{\Phi(t_j - t_{j-1}, Y_{[t_{j-1}, t_j]}) + f(Y_{t_{j-1}}, \sigma_{k-1}^2) - \frac{d}{2}\sigma_{k-1}^2\right\}}.$$

(c) Update μ and M : sample $\tilde{\mu}$ from $q(\mu|\mu_{k-1})$, compute the corresponding $\tilde{M} = \mathfrak{h}^{-1}(\tilde{\mu})$ and accept $\tilde{\mu}$, \tilde{M} with probability

$$\alpha^{(M)} = \frac{\pi_0^\mu(\tilde{\mu})}{\pi_0^\mu(\mu_{k-1})} \exp\left\{\sum_{j=1}^N \left[\Phi(t_j - t_{j-1}, \tilde{Y}_{[t_{j-1}, t_j]}) - \Phi(t_j - t_{j-1}, Y_{[t_{j-1}, t_j]})\right]\right\},$$

where we obtain $\tilde{Y}_{[t_{j-1}, t_j]}$ from $\{B_j^{(k)}, \theta_{k-1}, \tilde{M}, \sigma_k^2, \mathfrak{D}\}$ and $Y_{[t_{j-1}, t_j]}$ from $\{B_j^{(k)}, \theta_{k-1}, M_{k-1}, \sigma_k^2, \mathfrak{D}\}$ using equation (14) to approximately solve the SDE (18).

(d) Update θ similarly as μ .

Remark 2 (Time change). Since Φ in equation (22) explodes as $t \uparrow T$, Schauer et al. [2017] suggests time change and scaling to reduce the required number of imputed data points. Scaling will not be as effective here as in the univariate setting because it can only fit one of the directions involved, so the effect is less pronounced than in the univariate setting and we expect further lessening as dimension increases. Nonetheless, in this work, we adopt one time change function from Schauer et al. [2017] which maps s to $s(2 - s/T)$.

5. SIMULATION STUDY ON $\mathcal{SP}(2)$

5.1. Brownian bridges

We perform a simulation exercise to illustrate our proposed bridge sampling of Algorithm 1 and to compare the performance for the Euclidean, LE and AI metrics. The simulation scenario involves

sampling a standard Brownian bridge W_t conditioned on $\{W_0 = U, W_T = V\}$, $T = 0.1$, $n = 2$, in two cases: (i) with U and V lying far away from the boundary, $U = \begin{pmatrix} 2 & 1 \\ 1 & 2 \end{pmatrix}$, and $V = \begin{pmatrix} 3 & 1 \\ 1 & 2 \end{pmatrix}$ and (ii) with U and V lying close to the boundary, $U = \begin{pmatrix} 2 & 1.999 \\ 1.999 & 2 \end{pmatrix}$, $V = \begin{pmatrix} 3 & 2.435 \\ 2.435 & 2 \end{pmatrix}$. For the Euclidean metric we simply embed $\mathcal{SP}(n)$ in $\mathcal{S}(n)$ endowed with the Frobenius inner product and thus the Brownian motion in this case is the solution of the SDE $d\nu(W_t) = dB_t$ ($W_0 = U$), where B_t is the standard Brownian motion on \mathbb{R}^d and $\nu(W_t) \in \mathbb{R}^d$ contains only independent entries of W_t . The LE and AI metrics are based on Theorem 1 and Algorithm 1 respectively.

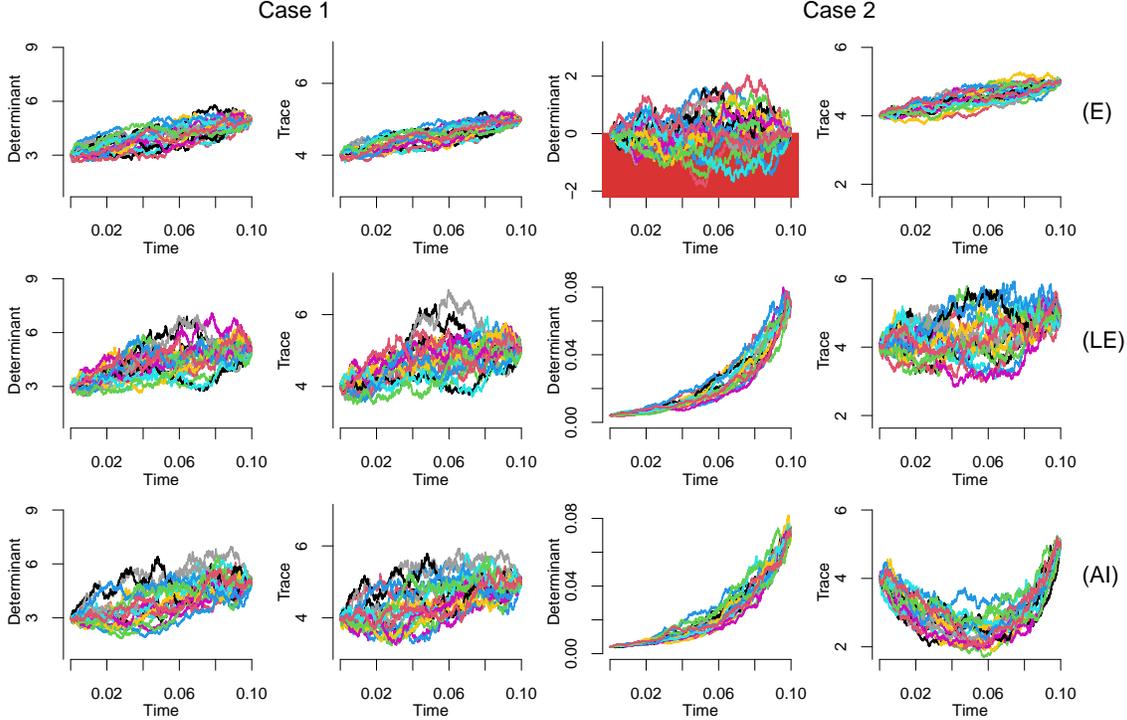


Figure 2: Time series for determinant and trace of 20 simulated Brownian bridges on $\mathcal{SP}(2)$ endowed with three different metrics (Euclidean: top row, log Euclidean: middle row, affine invariant: bottom row) in the 2 cases (i) and (ii). Red background area shows failure to be positive definite.

First we evaluate how Algorithm 1 performs compared with the naive simulation approach. For case (i) we obtain samples by forward-simulating the guided proposal in equation (18) and accepting with probability $\alpha^{(B)}$ in Algorithm 1. We then compare these bridges with the so-called true bridges, which are generated from the naive simulation approach by forward-solving the Riemannian Brownian motion X_t and picking only those paths X_t that satisfy $d_{\text{AI}}(X_{0.1}, V) < \epsilon$ for some $\epsilon > 0$. For different values of ϵ and number of imputed points m , we collect 1000 sampled bridges and carry out a Kolmogorov-Smirnov (K-S) test to compare the distribution of the true and approximated bridges at $t = T/2 = 0.05$; The Q-Q plots and the K-S p-values shown in Figure 3 indicate that the values $m = 2000$, $\epsilon = 0.05$ provide a good approximation and as $\epsilon \downarrow 0$ the two distributions get closer to each other. Note that for case (ii) we had a difficulty in sampling true bridges because V is too close to the boundary of $\mathcal{SP}(2)$ and points on the boundary are at infinite distance to covariance matrices. This difficulty escalates as dimension n increases. Beside high dependence with the diffusivity, the naive approach of simulating bridges has very low acceptance rate when the conditioned observation is close to being non-positive definite, thus it is clearly not a good approach to use in the data-augmentation algorithm.

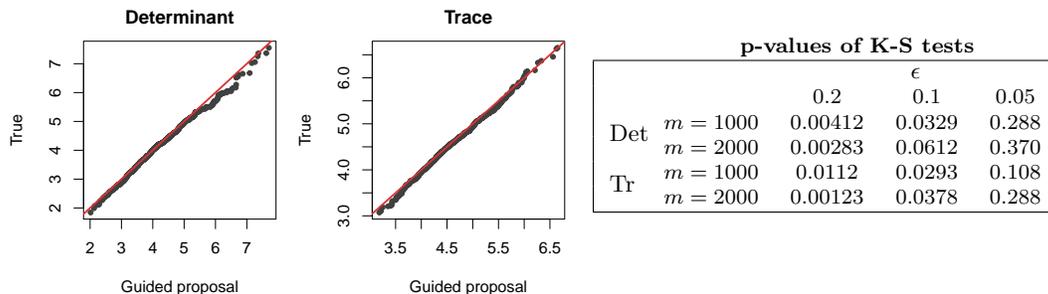


Figure 3: Comparison in distribution at $T/2$ of the true bridges and the guided proposal bridges, given that U and V are chosen as in Case 1. Right table: p-values from the KolmogorovSmirnov (K-S) tests, where ϵ and m are varied. Left figures: the Q-Q plots for determinant and trace when $m = 2000$ and $\epsilon = 0.05$.

Next, we illustrate the problems that arise when neglecting the geometric structure of $SP(n)$ as stated in Section 2.2.. Figure 2 depicts traces and determinants of 20 simulated bridges with each metric. First note that the Euclidean metric samples do not achieve positive definiteness and have a clear presence of the swelling effect. The LE and AI metrics sample points that lie in $\mathcal{SP}(n)$ while the determinants behave reasonably with respect to the conditional points. In fact, one can show following the lines of the case of geometric means in Arsigny et al. [2007] that the distribution of the determinant is the same for the LE and the AI metrics. The AI metric leads to less anisotropy that becomes more noticeable when the conditional points have eigenvalues close to zero, see plots about the trace in Fig. 2, indicating that the selection between LE and AI depends on the desirable property that one wishes to achieve.

5.2. Parameter estimation for the Affine-Invariant metric

We simulate $10^6 + 1$ equidistant time points X_t of the OU process in the case of the AI metric on $[0, 100]$ using equation (14) with model parameters $\theta = 0.5$, $M = \begin{pmatrix} 1 & 0.9 \\ 0.9 & 1 \end{pmatrix}$, $\sigma^2 = 1$ and $X_0 = I_2$ and take sub-samples at time points $\{0, 0.2, \dots, 100\}$. We apply the Algorithm 1 with time change assuming the prior distributions $\log \sigma^2$, $\log \theta \sim N(0, 4)$ and $\mu \sim N(0, 4 \times I_3)$.

Figure 4 is based on 1,000 burn-in and 4,000 MCMC iterations; it indicates that increasing m does not affect the mixing of the chain and improves, in some cases, the approximation to the marginal densities. We then run a longer MCMC chain of 50000 iterations while varying the value of m over 10, 50, 100 and 200. These chains are thinned out after a burn-in period of 2000 iterations and samples of 4000 points are collected from the target distributions. Figure 5 shows that the kernel density estimations of the marginal posterior distributions of $\{\theta, \mu, \sigma^2\}$ are approximately the same for $m = 100, 200$. Thus, $m = 100$ is considered to provide a sufficiently fine discretization for these data. The average proportions of accepting the bridges after the burn-in period are 72.7%, 70.2%, 68.5% and 67.3% for $m = 10, 50, 100$ and 200 respectively.

Our final investigation for Algorithm 1 is the prior reproduction test of Cook et al. [2006] to validate the Algorithm 1. We assume proper priors for $\{\theta, \mu, \sigma^2\} : \log \theta, \log \sigma^2 \sim N(0, 0.3^2)$, $\mu \sim N(0, 0.2 \times I_3)$ and $n = 500$. We generate, in turn, 1000 samples from the prior distributions and then, conditional on each sampled parameter vector, high-frequency observations on $[0, 50]$ at $5 \times 10^5 + 1$ equidistant time points and keep sub-samples at time points $\{0, 0.1, \dots, 50\}$. For the 1000 generated datasets we estimate the corresponding posterior densities using Algorithm 1 and test whether they come from the same distribution as the prior as this validates that our algorithm works properly. Figure 6 illustrates that the prior has been successfully replicated while, as expected, the parameters approximation improves as m increases.

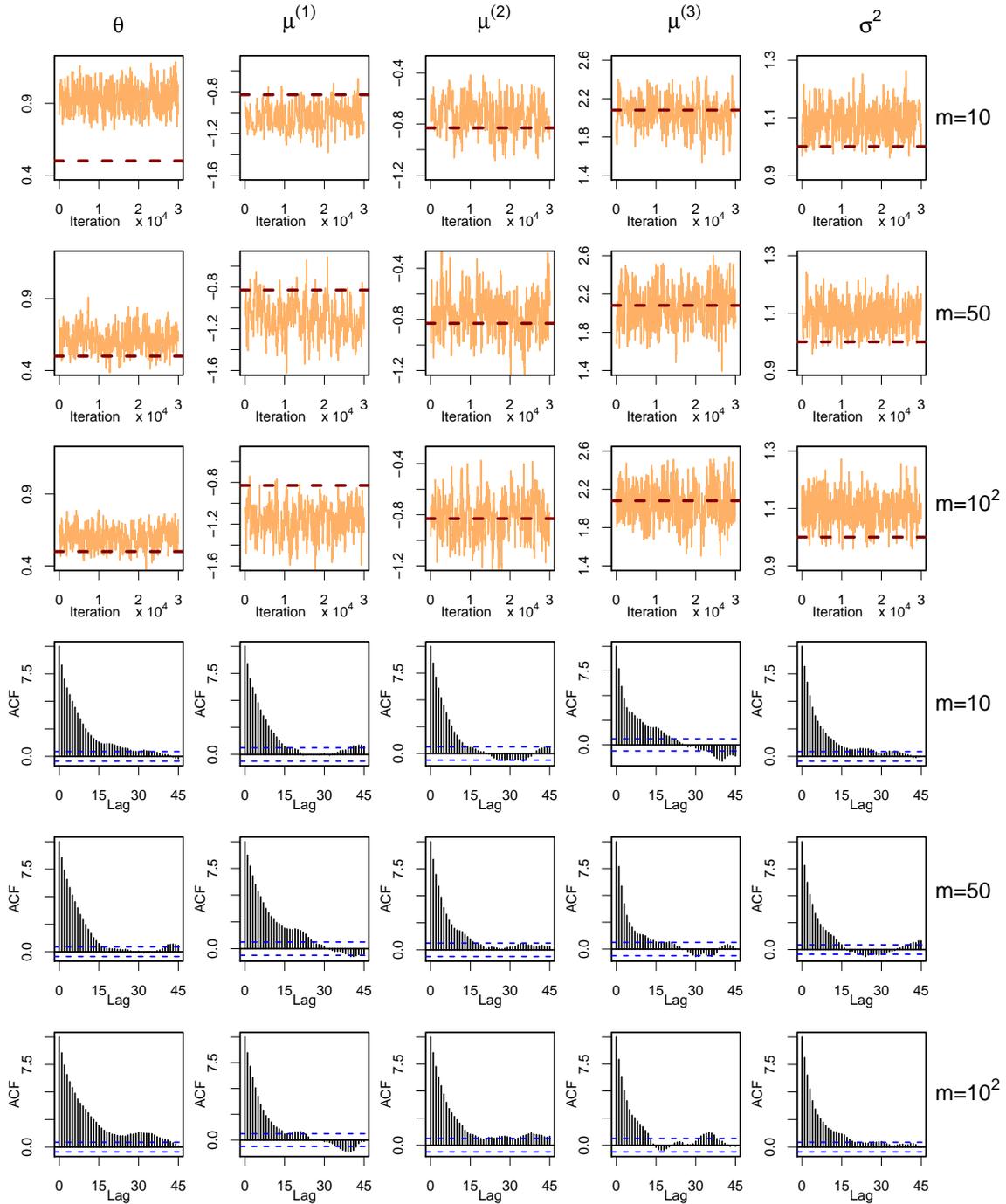


Figure 4: (Simulation study). m is varied over 10, 50 and 100. Top three rows: Traceplots from 4×10^3 iterates after discarding 10^3 iterations of burn-in, where true values are indicated with the red dashed lines. Bottom three rows: ACF plots based on these MCMC chains.

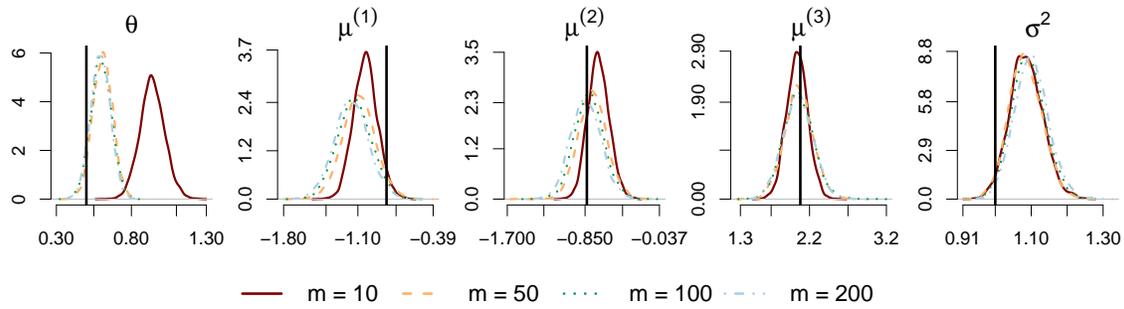


Figure 5: (Simulation study). Estimated posterior distribution of $\{\theta, \mu, \sigma^2\}$ using 5×10^4 MCMC iterations (2×10^3 burn-in discarded, thinned by 12), and the result for $M = \mathfrak{h}^{-1}(\mu)$ is given in Figure 11 (Appendix). True values are indicated by solid vertical black lines.

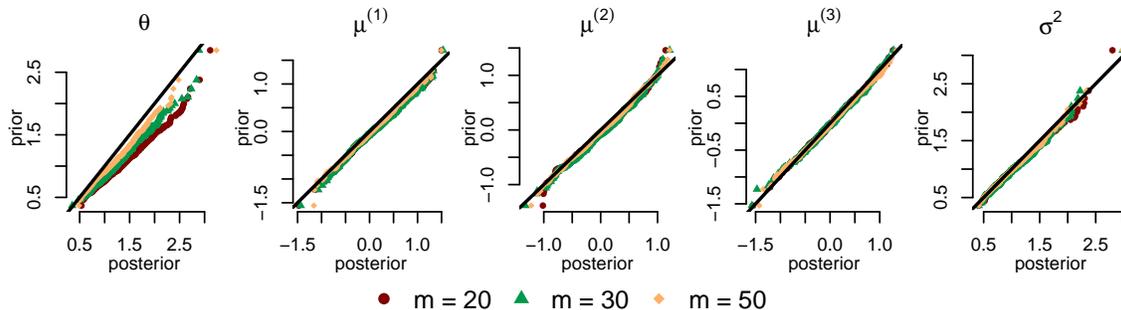


Figure 6: Prior reproduction test to validate the Algorithm 1 with time change: Q-Q plots of priors against posteriors from three model parameters.

6. APPLICATION TO FINANCE

6.1. Introduction

The widespread availability of intra-day high-frequency prices of financial assets has enabled the computation of consistent estimates of daily covariation of asset prices called realized covariances, introduced by Andersen et al. [2001a,b] and studied in detail by Barndorff-Nielsen and Shephard [2004]. The literature in discrete time series involves approaches based on Wishart-based distributions, see Asai and So [2014], Golosnoy et al. [2012], Gouriéroux et al. [2009], Jin and Maheu [2013], Yu et al. [2017], Jin et al. [2019]; matrix decompositions to deal with the positive definiteness requirement of the elements of the covariance matrices, see Bauer and Vorkink [2011], Chiriac and Voev [2011]; or ideas borrowed from the literature of multivariate GARCH models, see Noureldin et al. [2012], Hansen et al. [2014]. There has been quite a lot of evidence that direct modelling of realized covariances provides more precise forecasts than GARCH and stochastic volatility multivariate models that assume that the covariance matrices are unobserved latent matrices; see Golosnoy et al. [2012], Gouriéroux et al. [2009], Chiriac and Voev [2011], Bauer and Vorkink [2011], Noureldin et al. [2012], Golosnoy et al. [2012].

One important practical question in this framework is how covariance matrices vary at different time scales. It is well known work by Epps [1979] that correlations between assets decrease with the duration of investment horizons and this necessitates models that are frequency independent. Modelling realized covariances with diffusions offers a critical advantage over discrete time models because they allow inference of implied model dynamics and properties as well as forecasting at various frequencies that may differ from the observed data frequency. Moreover, continuous time

models are useful in irregularly spaced observed data and provide a clear advantage when used in pricing derivative instruments.

6.2. Data preprocessing

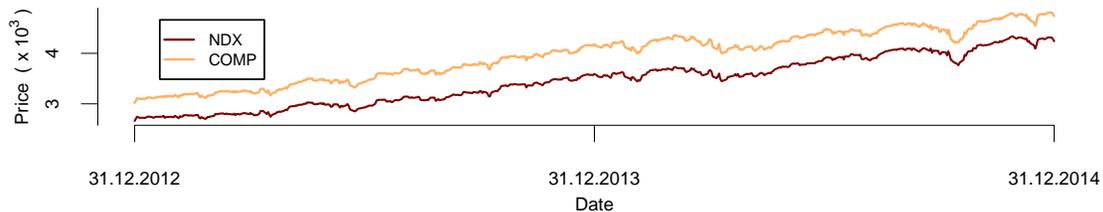


Figure 7: Time series of the closing price from two indices: NASDAQ Composite (COMP) and NASDAQ 100 (NDX) at the end of each working day from 31.12.2012 to 31.12.2014.

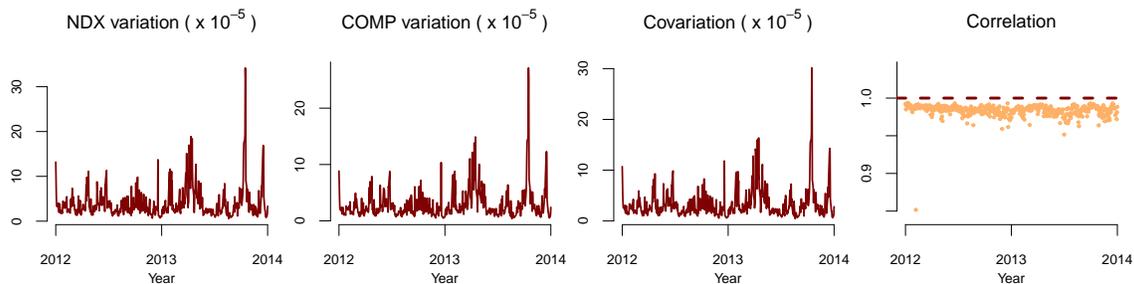


Figure 8: Time series of estimated covariance matrices based on the price of indices NDX and COMP. Estimated correlations are also plotted against time (red dashed line indicates correlation 1).

We estimate 2×2 daily covariance (volatility) matrices from NASDAQ Composite (COMP) and NASDAQ 100 (NDX) indices with data of 504 working days obtained from FirstRate Data [2019]; see Figure 7. For the estimation we used quadratic variation/covariation in the logarithm domain and assumed that the microstructure noise does not impact our estimates because the indices are very liquid, see Zhang [2011]. We verified this assumption by noting that estimates based on 5-minute inter-observation intervals are very similar. The pattern of the time series from the entries of covariance matrices in Figure 8 indicates a mean-reverting tendency and many observations lying close to the boundary of $\mathcal{SP}(2)$, making the Euclidean metric inappropriate and the Riemannian structures suitable.

6.3. Model fitting

Our time series has unevenly spaced observations due to weekends and holidays, so the imputed points m_j between the $(j - 1)$ th and j th consecutive observations are carefully chosen such that $\delta_t = (t_j - t_{j-1})/m_j$ is constant. We choose vague proper priors $\log \theta \sim N(0, 4)$, $\log \sigma^2 \sim N(-1, 4)$ and $\mu \sim N\left(\begin{pmatrix} -12 & -12 & 3 \end{pmatrix}^T, 4I_3\right)$. We adopt the algorithm by Roberts and Stramer [2001] in the case of the LE metric and Algorithm 1 with time change for the case of the AI metric. MCMC samples based on 10^5 iterations with varying values of δ_t were collected after a burn-in period of 4000 iterations and kernel density estimations are depicted in Figure 9.

Fewer imputed points are required for the LE metric than for the AI metric to adequately approximate the posterior densities which we attribute to different degrees of non-linearity: after transformation through the matrix logarithm, the LE problem is reduced to a linear problem whereas with

the AI metric, discretization including approximation of the horizontal lift takes place in the original domain. This domain can be seen to be less linear from the fact that covariance matrices are commutative under the logarithm product which is at the heart of the LE metric whereas exchanging the order of multiplication causes a different result for the AI metric. This difference is more pronounced near the boundary of the cone which is where the majority of our observations lie.

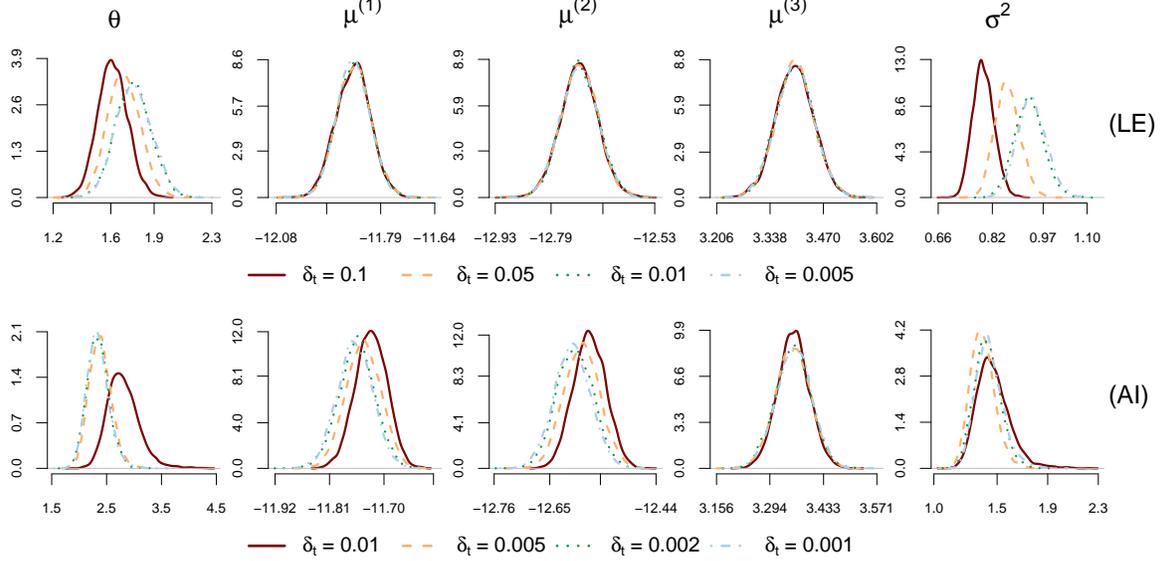


Figure 9: Estimated posterior distribution for $\{\theta, \mu, \sigma^2\}$ from financial data example using either the LE (top row) or the AI metric (bottom row), based on 10^5 MCMC iterations (4×10^3 burn-in discarded, thinned by 19). Moreover, the result for $M = \mathfrak{h}^{-1}(\mu)$ is given in Figure 13 (Appendix).

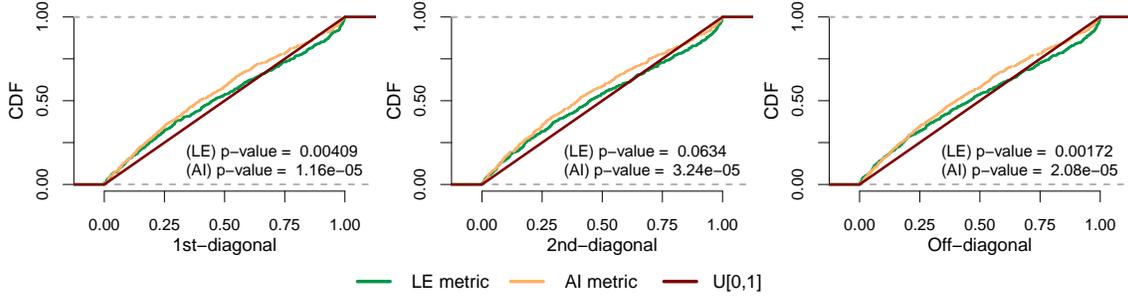


Figure 10: Empirical cumulative distribution of the generalized residuals for the entries from volatilities under two models using either the LE and the AI metric, in which model parameter estimates are the posterior mean. Moreover, p-values are obtained from the K-S tests of comparison to $U[0, 1]$ (red line).

We test the fit of the two models using the transition density-based approach by Hong and Li [2005]. For each model, we choose the posterior mean to estimate $\{\theta, M, \sigma^2\}$ and compute the generalized residuals $Z_j^{(i)}$, for $1 \leq j \leq 504$, as

$$Z_j^{(i)} = \int_{-\infty}^{y_j^{(i)}} p^{(i)}(t_j, v | t_{j-1}, y_{j-1}) dv. \quad (23)$$

where $\{y_j^{(i)}\}_{j=0, i=1}^{504, 3}$ denote the observations, $p^{(i)}$ are the marginal transition densities and $i = 1, 2, 3$ denotes the two diagonal and the off-diagonal entries respectively. The integral in equation (23) is

estimated by simulating $k = 3000$ points at t_j via equation (14) starting from y_{j-1} at time t_{j-1} . Under the null hypothesis that the observations come from the model, the realized generalized residuals $\{Z_j^{(i)}\}_{j=0}^{504}$ are i.i.d and follow the standard uniform distribution $U[0, 1]$ for all $i \in \{1, 2, 3\}$. The empirical cumulative distribution of these $Z_j^{(i)}$ are shown in Figure 10 with p-values from the Kolmogorov-Smirnov (K-S) test when comparing to $U[0, 1]$.

Figure 10 clearly indicates that the model using the LE metric fits the data better than the one using the AI metric for this particular dataset.

7. DISCUSSION

In summary, it is clear that the Euclidean metric, despite its simplicity, should not be used on $\mathcal{SP}(n)$. We instead suggest using either the LE or the AI metric. While both metrics behave similarly, the difference escalates when moving toward the boundary of $\mathcal{SP}(n)$ and the demonstrated goodness of fit testing can aid model choice. Although the AI metric generally has increased computational cost over the LE metric and exhibits slightly worse fit for our financial data example, it provides an alternative diffusion process on $\mathcal{SP}(n)$ with available Bayesian estimation. Moreover, the AI metric leads to less anisotropy than the LE metric, which is desirable in some other application areas, e.g. diffusion tensor imaging.

Cartan-Hadamard manifolds are diffeomorphic to Euclidean space, and the diffeomorphism maps can be obtained from the exponential map at any point, see Carmo [1992], Jost and Jost [2008]. Furthermore, following Karcher [1977], we note that $\nabla_{X_t} d^2$ can be expressed in terms of the logarithm map for more general manifolds so that our choice of drift is available for even more general Riemannian manifolds. Thus, our approach of sampling diffusion bridges and our proposed class of OU processes can be extended to any Cartan-Hadamard manifold on which there exists a suitable diffusion process with an explicit transition density function, e.g. hyperbolic spaces [Matsumoto, 2001, Nagano et al., 2019]. This opens up potential applications to phylogenetic trees [Nye et al., 2011] and electronic engineering [Huckemann et al., 2010] where this or closely related geometries are used. Extension to volatilities observed indirectly through the price processes generalizing the univariate models in Beskos et al. [2013] seems both feasible and practically relevant.

ACKNOWLEDGEMENT

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SUPPLEMENTARY MATERIAL

Supplementary material available online includes the proofs of Lemma 1, Theorems 1–2 and Propositions 1–2. Moreover, we include calculation of the approximation for ζ, ϕ and Φ , as well as prove the stochastic completeness on $\mathcal{SP}(n)$ with LE and AI metrics.

APPENDIX

We include the kernel density estimations for the marginal posterior distribution of the model parameters $\{\theta, M, \sigma^2\}$ in the simulation study (Fig. 11) and the application in finance (Fig. 13), which correspond to the estimated posterior distribution for $\{\theta, \mu, \sigma^2\}$ in Fig. 5 and Fig. 9, respectively. Furthermore, trace plots and ACF plots when using either the LE metric with $\delta_t = 0.01$ or the AI metric with $\delta_t = 0.001$ are shown in Figure 12.

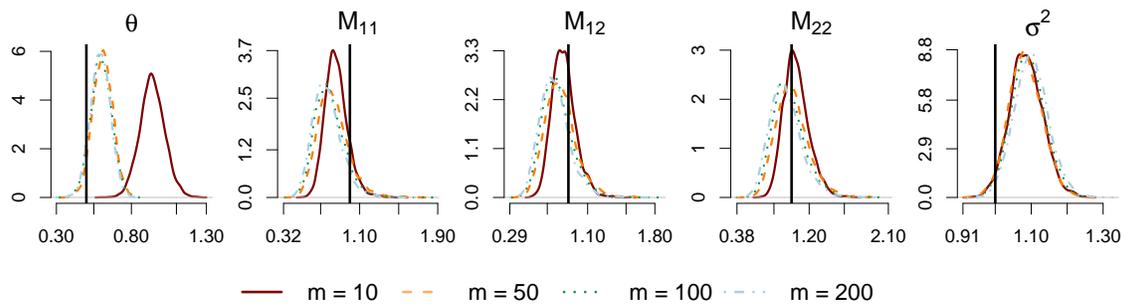


Figure 11: (Simulation study). Estimated posterior distribution of $\{\theta, M, \sigma^2\}$ using 5×10^4 MCMC iterations (2×10^3 burn-in discarded, thinned by 12). True values are indicated by solid vertical black lines.

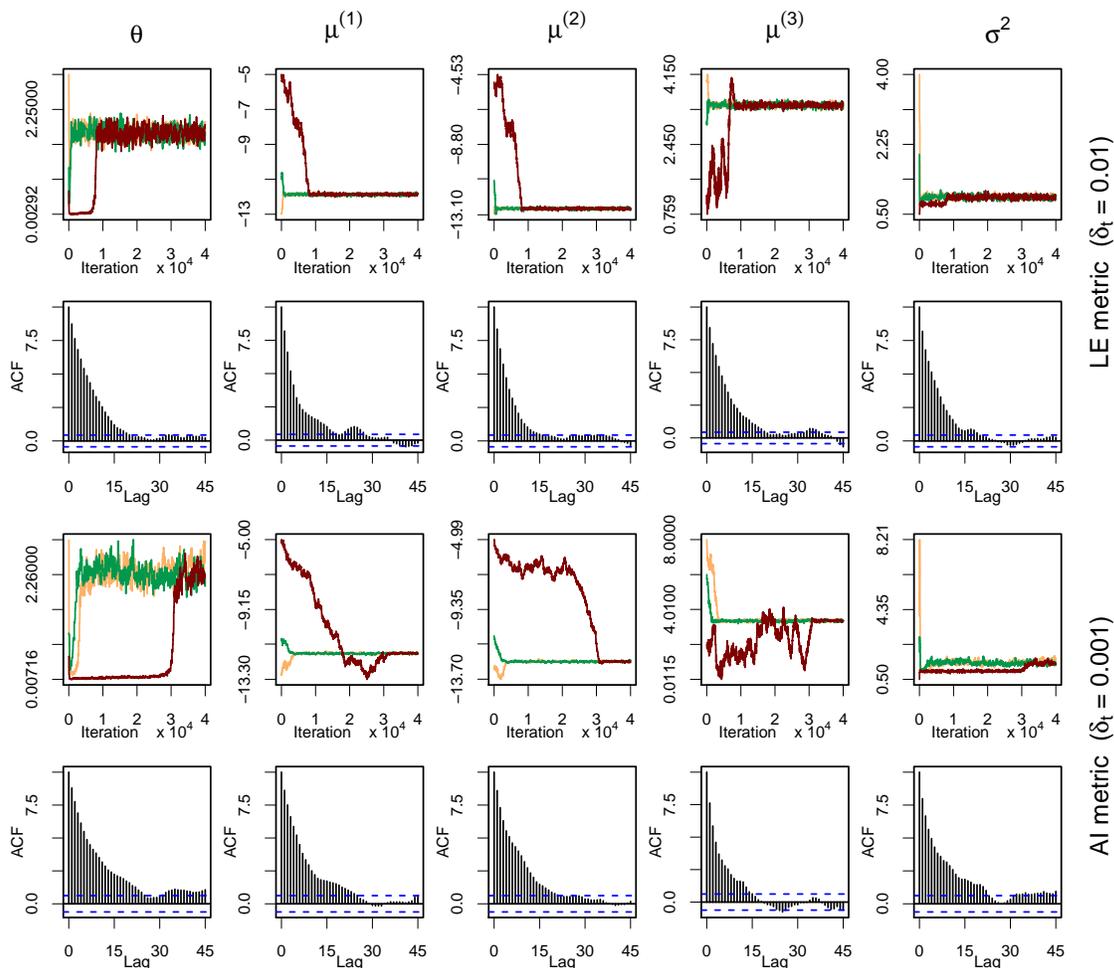


Figure 12: (Application in finance). MCMC trace plots of 4000 iterations using different starting points (orange, green and red) and ACF plots based on iterates 1000 – 4000 of the green chain with $\delta_t = 0.01$ in the case of the LE metric and $\delta_t = 0.001$ in the case of the AI metric .

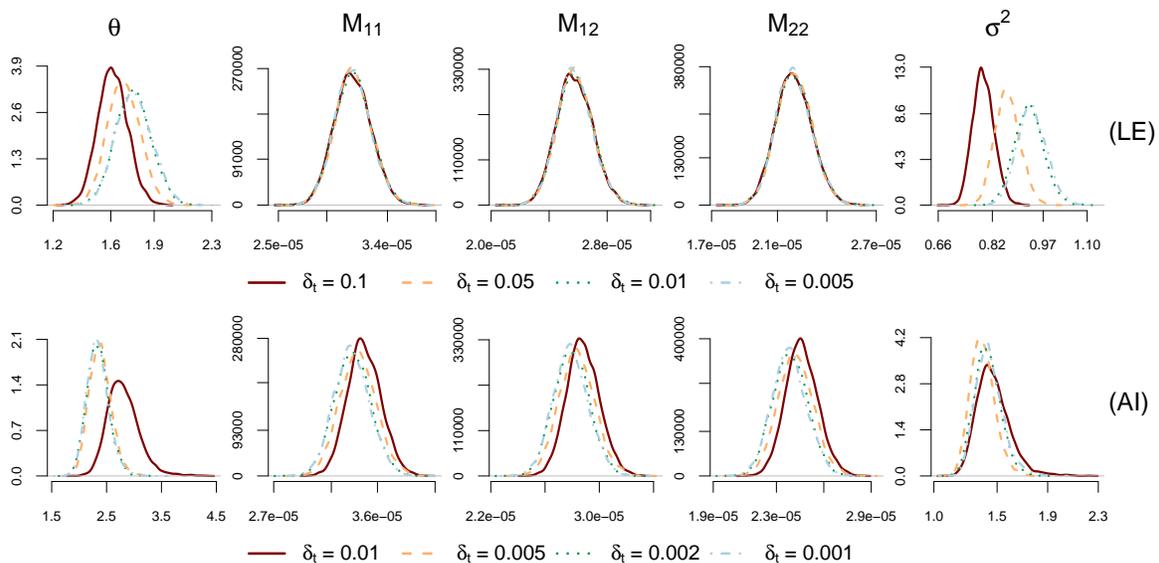


Figure 13: (Application in finance). Estimated posterior distribution for $\{\theta, M, \sigma^2\}$ using either the LE metric (top row) or the AI metric (bottom row), based on 10^5 MCMC iterations (4×10^3 burn-in discarded, thinned by 19).

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Supplementary material for “Inference for partially observed Riemannian Ornstein–Uhlenbeck diffusions of covariance matrices”

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Abstract

This supplement is divided into three sections. Section 1. contains the proof of Theorem S1, while Section 3. discusses stochastic completeness of $\mathcal{SP}(n)$ when equipping either the LE and the AI metric. Finally, we study the absolute continuity of the measure coming from the guided proposal on $\mathcal{SP}(n)$ equipped with the AI metric in Section 2.. We provide the proofs of Lemma S1, Theorem S2, Propositions S2–S3, together with the calculation of the approximation for ζ, ϕ and Φ in this section.

1. STOCHASTIC DIFFERENTIAL EQUATIONS ON $\mathcal{SP}(n)$ EQUIPPED WITH THE LE METRIC

Theorem S1. *Consider the following SDE on $\mathcal{SP}(n)$ endowed with the LE metric :*

$$dX_t = V(X_t) dt + F_{X_t}(a(X_t) dB_t) \quad (X_0 = P), \quad (\text{S1})$$

where V is a smooth vector field with smooth function $a : \mathcal{SP}(n) \rightarrow \mathbb{R}^{d \times d}$, B_t is \mathbb{R}^d -valued Brownian Motion and the function F is as defined in the SDE of OU processes on $\mathcal{SP}(n)$. Then the problem of solving the SDE (S1) on $\mathcal{SP}(n)$ is the same as solving the following SDE on \mathbb{R}^d :

$$dx_t = v(x_t) dt + \tilde{a}(x_t) dB_t \quad (x_0 = p), \quad (\text{S2})$$

(i) $p = \mathfrak{h}(P)$, $x_t = \mathfrak{h}(X_t)$ for all $t > 0$ and $\tilde{a} : \mathbb{R}^d \rightarrow \mathbb{R}^{d \times d}$, $e \mapsto a(\mathfrak{h}^{-1}(e))$,

(ii) Function $v = (v^{(j)})$ with $v^{(j)} \in C^\infty(\mathbb{R}^d)$, $x_t \mapsto \langle (d_{X_t} \log)V(X_t), S_j \rangle_F$ ($1 \leq j \leq d$).

Proof. As V is a smooth vector field on $\mathcal{SP}(n)$, there always exist functions $f^{(j)} \in C^\infty(\mathcal{SP}(n))$, $1 \leq j \leq d$ such that

$$V(X_t) = \sum_{j=1}^d f^{(j)}(X_t) E_j^{\text{LE}}(X_t) = \sum_{j=1}^d f^{(j)}(X_t) \{(d_{X_t} \log)^{-1} S_j\}.$$

Hence, $(d_{X_t} \log)V(X_t) = \sum_{j=1}^d f^{(j)}(X_t) S_j$, which are smooth on $\mathcal{S}(n)$ with respect to X_t as the function \log is diffeomorphic on $\mathcal{SP}(n)$. Since the $\{S_i\}_{i=1}^d$ are orthonormal, $f^{(j)}(X_t) = \langle (d_{X_t} \log)V(X_t), S_j \rangle_F$, which implies $v^{(j)}(x_t) = f^{(j)}(\mathfrak{h}^{-1}(x_t))$. Furthermore, since \mathfrak{h}^{-1} and f are both smooth, the function v is smooth on \mathbb{R}^d .

On the other hand, for an arbitrary $e = \sum_{j=1}^d \epsilon_j e_j \in \mathbb{R}^d$:

$$F_{X_t}(e) = \sum_{j=1}^d \epsilon_j E_j^{\text{LE}}(X_t) = (d_{X_t} \log)^{-1} \left(\sum_{j=1}^d \epsilon_j S_j \right) \iff (d_{X_t} \log) F_{X_t}(e) = \sum_{j=1}^d \epsilon_j S_j$$

Applying the push-forward by the matrix logarithm function (i.e. $d_{X_t} \log$) with the fact that $\log : \mathcal{SP}(n) \rightarrow \mathcal{S}(n)$ is a global diffeomorphism in using the result in Hsu [2002, Proposition 1.2.4, Page 20], the diffusion process x_t satisfies

$$\sum_{j=1}^d dx_t^{(j)} S_j = \sum_{j=1}^d v^{(j)}(x_t) S_j dt + \sum_{i,j=1}^d a_{ij}(x_t) S_j dB_t^i$$

The second term of the Ito formulation vanishes because $\mathcal{SP}(n)$ equipped with the LE metric has null sectional curvature everywhere. Removing the standard symmetric basis $\mathfrak{B}_d = \{S_i\}_{i=1}^d$, we get the desired result. \square

2. ABSOLUTE CONTINUITY OF THE GUIDED PROPOSAL ON $\mathcal{SP}(n)$ EQUIPPED WITH THE AFFINE-INVARIANT METRIC

Let us fix the standard basis $\{e_i\}_{i=1}^d$ for \mathbb{R}^d and consider the problem of X_t being the OU process on $\mathcal{SP}(n)$ equipped with the AI metric with its law \mathbb{P}_t :

$$dX_t = -\frac{\theta}{2} \nabla_{X_t} \{d^2(X_t, M)\} dt + F_{X_t}(\sigma dB_t) \quad (X_0 = P). \quad (\text{S3})$$

Our simulation method involves the Euler–Maruyama approximation via the exponential map,

$$X_{t+\delta_t} = \text{Exp}_{X_t} \left\{ -\frac{\theta}{2} \nabla_{X_t} \{d^2(X_t, M)\} \delta_t + \sum_{j=1}^d (B_{t+\delta_t}^j - B_t^j) \sigma E_j(X_t) \right\} \text{ for } \delta_t > 0. \quad (\text{S4})$$

We want to sample from the target diffusion bridge $X_t^* = \{X_t, 0 \leq t \leq T \mid X_0 = U, X_T = V\}$ with its corresponding law \mathbb{P}_t^* by introducing the guided proposal X_t^\diamond , which is the solution of the following SDE with its law \mathbb{P}_t^\diamond :

$$dX_t^\diamond = \left(\theta \text{Log}_{X_t^\diamond}^{\text{AI}} M + \frac{\text{Log}_{X_t^\diamond}^{\text{AI}} V}{T-t} \right) dt + F(X_t^\diamond)(\sigma dB_t) \quad (X_0^\diamond = U). \quad (\text{S5})$$

We aim to show the absolute continuity among three measures \mathbb{P}_t , \mathbb{P}_t^* and \mathbb{P}_t^\diamond in Theorem S2 below. In preparation for this main theorem, we will compute the Riemannian gradient of distance square in Proposition S2 and construct the stochastic development of smooth curves in Proposition S3 in the case of the LE and the AI metrics. Moreover, since the expression of the horizontal lift in local coordinates in the case of the AI metric does not admit explicit formulae, we approximate them in Corollary S1 and provide the calculation of Christoffel symbol in Lemma S1.

Proposition S2 (Riemannian gradient of distance squared).

(i) (LE metric). The set $\mathfrak{B}_d^{\text{LE}} = \{E_i^{\text{LE}}\}_{i=1}^d$ is an orthonormal frame on the tangent bundle TM , where for any $P \in \mathcal{SP}(n)$:

$$E_i^{\text{LE}}(P) = d_{\log P} \exp S_i = (d_P \log)^{-1} S_i \quad \text{for } 1 \leq i \leq d. \quad (\text{S6})$$

Moreover, the Riemannian gradient of distance squared for any fixed point $Q \in \mathcal{SP}(n)$ is $\nabla_P \{d_{\text{LE}}^2(P, Q)\} = -2(d_P \log)^{-1}(\log Q - \log P) = -2 \text{Log}_P^{\text{LE}}(Q)$.

(ii) (AI metric). The set $\mathfrak{B}_d^{AI} = \{E_i^{AI}\}_{i=1}^d$ is an orthonormal frame on the tangent bundle TM , where for any $P \in \mathcal{SP}(n)$:

$$E_i^{AI}(P) = P^{1/2} \star S_i \quad \text{for } 1 \leq i \leq d. \quad (S7)$$

Moreover, the Riemannian gradient of distance squared for any fixed point $Q \in \mathcal{SP}(n)$ is $\nabla_P \{d_{AI}^2(P, Q)\} = -2 \sum_{i=1}^d \langle \log(P^{-1/2}QP^{-1/2}), S_i \rangle_F E_i^{AI}(P) = -2 \text{Log}_P^{AI}(Q)$.

Proof. (i) Notice that the function $\log : \mathcal{SP}(n) \rightarrow \mathcal{S}(n)$ is bijective, thus there always exists uniquely $Q_i \in \mathcal{SP}(n)$ such that $\log Q_i - \log P = S_i$ for all $1 \leq i \leq d$. And $E_i^{\text{LE}}(P)$ simply equals to $\text{Log}_P^{\text{LE}}(Q_i) \in T_P \mathcal{SP}(n)$. Moreover, since matrix exponential and logarithm are smooth on $\mathcal{SP}(n)$, $E_i^{\text{LE}}(P)$ are smooth for all $1 \leq i \leq d$. Moreover, for any $1 \leq i, j \leq d$, we have

$$g_P^{\text{LE}}(E_i^{\text{LE}}(P), E_j^{\text{LE}}(P)) = \langle d_P \log E_i^{\text{LE}}(P), d_P \log E_j^{\text{LE}}(P) \rangle_F = \langle S_i, S_j \rangle_F = \delta_{ij},$$

where δ_{ij} is the Kronecker delta. Thus, the first argument is proved.

We have:

$$\nabla_P \{d_{\text{LE}}^2(P, Q)\} = \sum_{i=1}^d df_P \{E_i^{\text{LE}}(P)\} E_i^{\text{LE}}(P),$$

where $f(P) = d_{\text{LE}}^2(P, Q) = \langle \log Q - \log P, \log Q - \log P \rangle_F$ and $df_P \{E_i^{\text{LE}}(P)\}$ is the differential of f at P in the direction $E_i^{\text{LE}}(P)$. Since the matrix logarithm is a diffeomorphism and g_P^{LE} is induced from the Frobenius inner product through the derivative of the matrix logarithm, we get

$$\begin{aligned} df_P \{E_i^{\text{LE}}(P)\} &= 2 g_P^{\text{LE}}(\nabla_{E_i(P)} \text{Log}_P^{\text{LE}} M, \text{Log}_P^{\text{LE}} Q) \\ &= 2 \langle d_P(\log Q - \log P) E_i(P), \log Q - \log P \rangle_F \\ &= 2 \langle -d_P \log E_i(P), \log Q - \log P \rangle_F \\ &= -2 \langle S_i, \log Q - \log P \rangle_F \\ \Rightarrow \nabla_P \{d_{\text{LE}}^2(P, Q)\} &= -2 (d_P \log)^{-1}(\log Q - \log P) = -2 \text{Log}_P^{\text{LE}}(Q) \end{aligned}$$

(ii) Since the function $\log : \mathcal{SP}(n) \rightarrow \mathcal{S}(n)$ is bijective, there always exists uniquely $Q_i \in \mathcal{SP}(n)$ such that $\log(P^{-1/2}Q_iP^{-1/2}) = S_i$ for all $1 \leq i \leq d$. And $E_i(P) = \text{Log}_P^{AI}(Q_i) \in T_P \mathcal{SP}(n)$. Moreover, for any $1 \leq i, j \leq d$, we have

$$\begin{aligned} g_P^{AI}(E_i^{AI}(P), E_j^{AI}(P)) &= \left\langle P^{-1/2} E_i^{AI}(P) P^{-1/2}, P^{-1/2} E_j^{AI}(P) P^{-1/2} \right\rangle_F \\ &= \langle S_i, S_j \rangle_F = \delta_{ij}, \end{aligned}$$

where δ_{ij} is the Kronecker delta. We hence obtain a global orthonormal basis field \mathfrak{B}_d^{AI} on $\mathcal{SP}(n)$, by the congruent transformation of $P^{1/2}$.

Similarly to the case of the LE metric, we have

$$\nabla_P \{d_{AI}^2(P, Q)\} = \sum_{i=1}^d df_P \{E_i^{AI}(P)\} E_i^{AI}(P),$$

where $f(P) = d_{AI}^2(P, M)$ and $df_P \{E_i^{AI}(P)\}$ is the differential of f at P in the direction $E_i^{AI}(P)$. Consider a geodesic curve $\gamma(t) = P^{1/2} \exp(t S_i) P^{1/2}$, which satisfies $\gamma(0) = P$, $\dot{\gamma}(0) = P^{1/2} S_i P^{1/2} = E_i^{AI}(P)$, and we have

$$df_P(E_i^{AI}(P)) = \left. \frac{d}{dt} f(\gamma(t)) \right|_{t=0},$$

$$f(\gamma(t)) = \|\log \{M^{-1/2}\gamma(t)Q^{-1/2}\}\|_F^2 = \text{tr} \left(\log^2 \{Q^{-1/2}\gamma(t)Q^{-1/2}\} \right).$$

Letting $\phi(t) = Q^{-1/2}\gamma(t)Q^{-1/2}$, then for any non-zero $v \in \mathbb{R}^n$,

$$v^T \phi(t)v = (P^{1/2}Q^{-1/2}v)^T \exp(tS_i)(P^{1/2}Q^{-1/2}v) > 0$$

as $\exp(tS_i) \in \mathcal{SP}(n)$, thus $\phi(t) \in \mathcal{SP}(n)$, $\forall t \geq 0$. We then can apply the result by Moakher [2005] to $\phi(t)$, which implies

$$\frac{d}{dt} \text{tr} (\log^2 \phi(t)) = 2 \text{tr} \left(\log \phi(t) \{\phi(t)\}^{-1} \frac{d}{dt} \phi(t) \right)$$

$$\begin{aligned} \Rightarrow \frac{d}{dt} f(\gamma(t)) \Big|_{t=0} &= \frac{d}{dt} \text{tr} (\log^2 \phi(t)) \Big|_{t=0} = 2 \text{tr} \left(\log \phi(0) \{\phi(0)\}^{-1} \dot{\phi}(0) \right) \\ &= 2 \text{tr} \left(\log \phi(0) \{\phi(0)\}^{-1} Q^{-1/2} \dot{\gamma}(0) Q^{-1/2} \right) \\ &= 2 \text{tr} \left(\log(Q^{-1/2} P Q^{-1/2}) Q^{-1/2} P^{1/2} S_i P^{1/2} Q^{-1/2} \right) \\ &= 2 \text{tr} \left(\log(P^{1/2} Q^{-1} P^{1/2}) S_i \right) = -2 \left\langle \log(P^{-1/2} Q P^{-1/2}), S_i \right\rangle_F \end{aligned}$$

Thus, $\nabla_P \{d_{\text{AI}}^2(P, Q)\} = -2 \sum_{i=1}^d \langle \log(P^{-1/2} Q P^{-1/2}), S_i \rangle_F E_i^{\text{AI}}(P) = -2 \text{Log}_P^{\text{AI}} Q$. \square

Lemma S1. *The Christoffel symbols at any $P \in \mathcal{SP}(n)$ with respect to the basis $\mathfrak{B}_d^{\text{AI}}$ are $\Gamma_{ij}^k(P) = \langle -(S_i S_j + S_j S_i), S_k \rangle_F / 2$, which are constant (i.e. they do not depend on P).*

Proof. We take the result by Moakher and Zérai [2011] that the Levi-Civita connection of $\mathcal{SP}(n)$ equipped with the AI metric are given as follows:

$$(\nabla_X Y)_P = -\frac{1}{2}(X_P P^{-1} Y_P + Y_P P^{-1} X_P) \quad \text{for } X, Y \in \Gamma(T\mathcal{SP}(n)).$$

Replacing $X = E_i^{\text{AI}}(P)$ and $Y = E_j^{\text{AI}}(P)$, we get

$$\begin{aligned} (\nabla_{E_i^{\text{AI}}} E_j^{\text{AI}})_P &= P^{1/2} \left\{ -\frac{1}{2}(S_i S_j + S_j S_i) \right\} P^{1/2} = \sum_{k=1}^d \left\langle -\frac{1}{2}(S_i S_j + S_j S_i), S_k \right\rangle_F E_k^{\text{AI}}(P) \\ &\Rightarrow \Gamma_{ij}^k(P) = \left\langle -\frac{1}{2}(S_i S_j + S_j S_i), S_k \right\rangle_F \end{aligned}$$

and clearly, Christoffel symbols do not depend on P . \square

Proposition S3 (Horizontal lift of smooth curves). *Suppose that $\gamma(t)$ is a smooth curve on $\mathcal{SP}(n)$ with $\gamma(0) = P$ and some smooth curve u_t such that $\pi(u_t) = \gamma(t)$ for all $t > 0$, where $\pi : \mathfrak{F}(\mathcal{SP}(n)) \rightarrow \mathcal{SP}(n)$ is the canonical projection map.*

- (i) (LE metric). *Suppose the expression of u_t in local coordinates $\{E_1^{\text{LE}}(\gamma(t)), \dots, E_d^{\text{LE}}(\gamma(t)), e_1, \dots, e_d\}$ is $u_t = (\gamma(t), \delta)$ with $\delta = (\delta_{ij})$, the Kronecker delta, then u_t is the unique horizontal lift of the curve $\gamma(t)$ from u_0 , where $u_0(e_i) = E_i^{\text{LE}}(P)$ for all $1 \leq i \leq d$.*
- (ii) (AI metric). *Suppose the expression of u_t in local coordinates $\{E_1^{\text{AI}}(\gamma(t)), \dots, E_d^{\text{AI}}(\gamma(t)), e_1, \dots, e_d\}$ is $u_t = (\gamma(t), \zeta)$ with $\zeta = (\zeta_j^i)$ and $\zeta_j^i : \mathcal{SP}(n) \rightarrow \mathbb{R}$ are continuous functions, then u_t is the*

unique horizontal lift of the curve $\gamma(t)$ from $u_0 : \mathbb{R}^d \rightarrow T_P \mathcal{SP}(n)$ with $u_0(e_i) = E_i^{AI}(P)$ ($1 \leq i \leq d$), i.e. $\zeta_j^i(P) = \delta_{ij}$ if and only if functions ζ_j^i exist uniquely and satisfy:

$$\sum_{r=1}^d \alpha_r(\gamma(t)) \left\{ (E_r^{AI} \zeta_j^i)_{\gamma(t)} + \sum_{k=1}^d \zeta_k^i(\gamma(t)) \Gamma_{rk}^j(\gamma(t)) \right\} = 0 \quad \text{for all } 1 \leq i, j \leq d, \quad (\text{S8})$$

where $\Gamma_{rk}^j(\gamma(t)) = \Gamma_{rk}^j$, the Christoffel symbols, were shown constant in Lemma S1 and $\dot{\gamma}(t) = \sum_{r=1}^d \alpha_r(\gamma(t)) E_r^{AI}(\gamma(t))$ for some functions $\alpha_r \in C^\infty(\mathcal{SP}(n))$.

Moreover, if $\alpha_r(\gamma(t)) \neq 0$ for only one $r \in \{1, \dots, d\}$, then functions $\zeta_j^i(\gamma(t))$ have to satisfy

$$(E_r^{AI} \zeta_j^i)_{\gamma(t)} = - \sum_{k=1}^d \zeta_k^i(\gamma(t)) \Gamma_{rk}^j \quad \text{for all } 1 \leq i, j \leq d.$$

Proof. (i) We are given $\pi(u_t) = \gamma(t)$ with fixed initial values $\gamma(0) = P$ and $u_0 \in \mathfrak{F}(\mathcal{SP}(n))_P$, thus it is sufficient to show that $u_t(e)$ is parallel along the curve $\gamma(t)$ for any $e \in \mathbb{R}^d$, i.e. $\nabla_{\dot{\gamma}(t)} u_t(e) = 0$.

Consider an arbitrary $e \in \mathbb{R}^d$, then $e = \sum_{i=1}^d \epsilon_i e_i$ for some $\epsilon_i \in \mathbb{R}$. By definition, $u_t(e) = \sum_{i=1}^d \epsilon_i E_i(X_t)$. On the other hand, we also can write $\dot{\gamma}(t) = \sum_{i=1}^d \alpha_i(\gamma(t)) E_i(\gamma(t))$ for some functions $\alpha_i \in C^\infty(\mathcal{SP}(n))$. Since ϵ_i does not depend on $\gamma(t)$ for all $1 \leq i \leq d$, we have :

$$\begin{aligned} \nabla_{\dot{\gamma}(t)} u_t(e) &= \nabla_{\{\sum_{i=1}^d \alpha_i(\gamma(t)) E_i(\gamma(t))\}} \left\{ \sum_{j=1}^d \epsilon_j E_j(\gamma(t)) \right\} \\ &= \sum_{i,j=1}^d \alpha_i(\gamma(t)) \epsilon_j \left\{ \nabla_{E_i(X_t)} E_j(\gamma(t)) \right\} = 0. \end{aligned}$$

Here we used the fact that $\nabla_{E_i^{LE}(Q)} E_j^{LE}(Q) = 0$ ($1 \leq i, j \leq d$) for all $Q \in \mathcal{SP}(n)$.

(ii) Firstly, let us suppose that there exist such functions ζ_j^i , which satisfy equation (S8). Since $\pi(u_t) = \gamma(t)$ with fixed u_0 and $\gamma(0) = P$, it is sufficient to show u_t is the unique horizontal lift of $\gamma(t)$ by showing for any $e \in \mathbb{R}^d$, $u_t(e)$ is parallel along the curve $\gamma(t)$, i.e. $\nabla_{\dot{\gamma}(t)} u_t(e) = 0$.

Consider an arbitrary $e \in \mathbb{R}^d$, i.e. $e = \sum_{i=1}^d \epsilon_i e_i$ for some scalar ϵ_i , then

$$u_t(e) = \sum_{i,j=1}^d \epsilon_i \zeta_j^i(\gamma(t)) E_j^{AI}(\gamma(t)).$$

Using the given condition in equation (S8), we get

$$\begin{aligned} \nabla_{\dot{\gamma}(t)} u_t(e) &= \sum_{r=1}^d \alpha_r(\gamma(t)) \left(\nabla_{E_r^{AI}(\gamma(t))} \left\{ \sum_{i,j=1}^d \epsilon_i \zeta_j^i(\gamma(t)) E_j^{AI}(\gamma(t)) \right\} \right) \\ &= \sum_{r=1}^d \alpha_r(\gamma(t)) \left(\sum_{i,j=1}^d \epsilon_i \nabla_{E_r^{AI}(\gamma(t))} \{ \zeta_j^i(\gamma(t)) E_j^{AI}(\gamma(t)) \} \right) \\ &= \sum_{r=1}^d \alpha_r(\gamma(t)) \left(\sum_{i,j=1}^d \epsilon_i \{ (E_r^{AI} \zeta_j^i)_{\gamma(t)} E_j^{AI}(\gamma(t)) + \zeta_j^i(X_t) \nabla_{E_r^{AI}(\gamma(t))} E_j^{AI}(\gamma(t)) \} \right) \\ &= \sum_{r,i,j=1}^d \alpha_r(\gamma(t)) \epsilon_i (E_r^{AI} \zeta_j^i)_{\gamma(t)} E_j^{AI}(\gamma(t)) \end{aligned}$$

$$\begin{aligned}
& + \sum_{r=1}^d \alpha_r(\gamma(t)) \sum_{i,j=1}^d e_i \zeta_j^i(\gamma(t)) \left(\sum_{k=1}^d \Gamma_{rk}^k(\gamma(t)) E_k^{\text{AI}}(\gamma(t)) \right) \\
& = \sum_{i,j=1}^d e_i \left(\sum_{r=1}^d \alpha_r(\gamma(t)) \left\{ (E_r^{\text{AI}} \zeta_j^i)_{\gamma(t)} + \sum_{k=1}^d \zeta_k^i(\gamma(t)) \Gamma_{rk}^j \right\} \right) E_j^{\text{AI}}(\gamma(t)) = 0.
\end{aligned}$$

On the other hand, if u_t is the horizontal lift of X_t starting from u_0 , for any $e \in \mathbb{R}^d$, $\nabla_{\dot{\gamma}(t)} u_t(e) = 0$.

$$\Rightarrow \sum_{r=1}^d \alpha_r(\gamma(t)) \left\{ (E_r^{\text{AI}} \zeta_j^i)_{\gamma(t)} + \sum_{k=1}^d \zeta_k^i(\gamma(t)) \Gamma_{rk}^j \right\} = 0 \quad \text{for all } 1 \leq i, j \leq d.$$

If only one $r \in \{1, \dots, d\}$ is such that $\alpha_r(X_t) \neq 0$, then clearly functions $\zeta_j^i(\gamma(t))$ have to satisfy

$$(E_r \zeta_j^i)_{\gamma(t)} = - \sum_{k=1}^d \zeta_k^i(\gamma(t)) \Gamma_{rk}^j \quad \text{for all } 1 \leq i, j \leq d$$

Uniqueness and existence of u_t result from the fact that ζ_j^i is the solution of a system of first order linear ordinary differential equations with initial conditions $\zeta_j^i(P) = \delta_{ij}$. \square

Theorem S2. For $t \in [0, T]$ the laws $\mathbb{P}_t, \mathbb{P}_t^\diamond$ and \mathbb{P}_t^* are absolutely continuous. Let $p(t, X_t; T, V)$ denote the true (unknown) transition density of moving from X_t at time t to V at time T and let $X_{[0:t]}^\diamond$ be the path of X_t^\diamond from time 0 to t , then the following Radon-Nikodym derivatives are given in closed form:

$$\frac{d\mathbb{P}_t}{d\mathbb{P}_t^\diamond}(X_{[0:t]}^\diamond) = \frac{\exp(f(X_0^\diamond))}{\exp(f(X_t^\diamond))} \exp \left\{ \Phi(t, X_{[0:t]}^\diamond) + \phi(t, X_{[0:t]}^\diamond) \right\}, \quad (\text{S9})$$

$$\frac{d\mathbb{P}_t^*}{d\mathbb{P}_t^\diamond}(X_{[0:t]}^\diamond) = \frac{p(t, X_t^\diamond; T, V)}{\exp(f(X_t^\diamond))} \frac{\exp(f(X_0^\diamond))}{p(0, U; T, V)} \exp \left\{ \Phi(t, X_{[0:t]}^\diamond) + \phi(t, X_{[0:t]}^\diamond) \right\}, \quad (\text{S10})$$

where f , ϕ and Φ are

$$f(X_t^\diamond) = -\frac{d_{AI}^2(X_t^\diamond, V)}{2\sigma^2(T-t)} = -\frac{\|\log((X_t^\diamond)^{-1/2} V (X_t^\diamond)^{-1/2})\|_F^2}{2\sigma^2(T-t)}, \quad (\text{S11})$$

$$\phi(t, X_{[0:t]}^\diamond) = \sum_{i,j=1}^d \int_0^t \frac{(\zeta_j^i(X_s^\diamond))^2}{2(T-s)} ds, \quad (\text{S12})$$

$$\begin{aligned}
\Phi(t, X_{[0:t]}^\diamond) &= \int_0^t \left[\theta g_{X_s^\diamond}^{\text{AI}} \left(\text{Log}_{X_s^\diamond}^{\text{AI}} M, \frac{\text{Log}_{X_s^\diamond}^{\text{AI}} V}{\sigma^2(T-s)} \right) \right. \\
&\quad \left. + \frac{1}{2} \sum_{i,j,r=1}^d g_{X_s^\diamond}^{\text{AI}} \left(\zeta_j^i(X_s^\diamond) \left(\sum_{l=1}^d \zeta_l^i(X_s^\diamond) \Gamma_{jl}^r + (E_j^{\text{AI}} \zeta_r^i)_{X_s^\diamond} \right) E_r^{\text{AI}}(X_s^\diamond), \frac{\text{Log}_{X_s^\diamond}^{\text{AI}} V}{T-s} \right) \right] ds,
\end{aligned} \quad (\text{S13})$$

with Γ_{jl}^r given in Lemma S1. The functions $\zeta = (\zeta_l^i)$ are the coefficients with respect to the basis $\mathfrak{B}_d^{\text{AI}}$ in the expression of the horizontal lift in local coordinates.

Proof. Using the Girsanov-Cameron-Martin Theorem in Elworthy [1982, Theorem 11C, Page 263], the measures \mathbb{P}_t and \mathbb{P}_t^\diamond are absolutely continuous, and the Radon-Nikodym derivative is given as

$$\frac{d\mathbb{P}_t}{d\mathbb{P}_t^\diamond}(X_{[0:t]}^\diamond) = \exp \left\{ - \int_0^t \frac{g_{X_s^\diamond}^{\text{AI}} \left(\text{Log}_{X_s^\diamond}^{\text{AI}} V, U_s^\diamond(dB_s) \right)}{\sigma(T-s)} - \frac{1}{2} \int_0^t \frac{\|\text{Log}_{X_s^\diamond}^{\text{AI}}\|_{g_{X_s^\diamond}^{\text{AI}}}^2}{\sigma^2(T-s)^2} ds \right\}, \quad (\text{S14})$$

where U_t^\diamond is the horizontal lift of the guided proposal X_t^\diamond .

On the other hand, we will apply Itô's formula from [Elworthy, 1982, Lemma 9B, Page 145] to the smooth function f while using some results:

$$\begin{aligned}
\nabla_{X_t^\diamond} f &= \frac{\text{Log}_{X_t^\diamond}^{\text{AI}} V}{\sigma^2(T-t)}, \\
\Delta f\{E_i^{\text{AI}}(X_t^\diamond), E_j^{\text{AI}}(X_t^\diamond)\} &= \frac{\delta_{ij}}{\sigma^2(T-t)}, \\
\frac{\partial f}{\partial t} &= -\frac{d_{\text{AI}}^2(V, X_t^\diamond)}{2\sigma^2(T-t)^2} = -\frac{\|\text{Log}_{X_t^\diamond}^{\text{AI}}\|_{g_{X_t^\diamond}}^2}{2\sigma^2(T-t)^2}, \\
U_t^\diamond(e_i) &= \sum_{j=1}^d \zeta_j^i(X_t^\diamond) E_j^{\text{AI}}(X_t^\diamond) \\
\Rightarrow \nabla_{U_t^\diamond(e_i)} U_t^\diamond(e_k) &= \sum_{j,l} \zeta_j^i(X_t^\diamond) \left(\sum_{r=1}^d \zeta_l^k(X_t^\diamond) \Gamma_{jl}^r E_r^{\text{AI}}(X_t^\diamond) + (E_j^{\text{AI}} \zeta_l^k)_{X_t^\diamond} E_l^{\text{AI}}(X_t^\diamond) \right).
\end{aligned}$$

Thus,

$$\begin{aligned}
f(X_t^\diamond) - f(X_0^\diamond) &= \int_0^t \frac{\partial f}{\partial s} ds + \int_0^t g_{X_s^\diamond}^{\text{AI}} \left(\theta \text{Log}_{X_s^\diamond}^{\text{AI}} M + \frac{\text{Log}_{X_s^\diamond}^{\text{AI}} V}{T-s}, \nabla_{X_s^\diamond} f \right) ds \\
&+ \int_0^t g_{X_s^\diamond}^{\text{AI}} (U_s^\diamond(\sigma dB_s), \nabla_{X_s^\diamond} f) + \frac{1}{2} \int_0^t g_{X_s^\diamond}^{\text{AI}} (\nabla_{U_s^\diamond(\sigma dB_s)} U_s^\diamond(\sigma dB_s), \nabla_{X_s^\diamond} f) \\
&+ \frac{1}{2} \int_0^t \Delta f\{U_s^\diamond(\sigma dB_s), U_s^\diamond(\sigma dB_s)\} \\
&= -\frac{1}{2} \int_0^t \frac{\|\text{Log}_{X_s^\diamond}^{\text{AI}}\|_{g_{X_s^\diamond}^{\text{AI}}}^2}{\sigma^2(T-s)^2} ds + \int_0^t \theta g_{X_s^\diamond}^{\text{AI}} \left(\text{Log}_{X_s^\diamond}^{\text{AI}} M, \frac{\text{Log}_{X_s^\diamond}^{\text{AI}} V}{\sigma^2(T-s)} \right) ds \\
&+ \int_0^t \frac{\|\text{Log}_{X_s^\diamond}^{\text{AI}}\|_{g_{X_s^\diamond}^{\text{AI}}}^2}{\sigma^2(T-s)^2} ds + \int_0^t g_{X_s^\diamond}^{\text{AI}} \left(\frac{\text{Log}_{X_s^\diamond}^{\text{AI}} V}{\sigma^2(T-s)}, U_s^\diamond(\sigma dB_s) \right) \\
&+ \frac{1}{2\sigma^2(T-s)} \int_0^t \sum_{i,k=1}^d d\langle B^i, B^k \rangle_s \left\{ \sum_{j,l=1}^d \delta_{jl} \zeta_j^i(X_s^\diamond) \zeta_l^k(X_s^\diamond) \sigma^2 \right. \\
&+ \left. \sum_{j,r=1}^d g_{X_s^\diamond}^{\text{AI}} \left(\sigma^2 \zeta_j^i(X_s^\diamond) \left(\sum_{l=1}^d \zeta_l^k(X_s^\diamond) \Gamma_{jr}^l + (E_j^{\text{AI}} \zeta_r^k)_{X_s^\diamond} \right) E_l^{\text{AI}}(X_s^\diamond), \text{Log}_{X_s^\diamond}^{\text{AI}} V \right) \right\} \\
&= \frac{1}{2} \int_0^t \frac{\|\text{Log}_{X_s^\diamond}^{\text{AI}}\|_{g_{X_s^\diamond}^{\text{AI}}}^2}{\sigma^2(T-s)^2} ds + \int_0^t \frac{g_{X_s^\diamond}^{\text{AI}} (\text{Log}_{X_s^\diamond}^{\text{AI}} V, U_s^\diamond(dB_s))}{\sigma(T-s)} \\
&+ \int_0^t \theta g_{X_s^\diamond}^{\text{AI}} \left(\text{Log}_{X_s^\diamond}^{\text{AI}} M, \frac{\text{Log}_{X_s^\diamond}^{\text{AI}} V}{\sigma^2(T-s)} \right) ds \\
&+ \frac{1}{2} \sum_{i,j,r=1}^d \int_0^t g_{X_s^\diamond}^{\text{AI}} \left(\zeta_j^i(X_s^\diamond) \left(\sum_{l=1}^d \zeta_l^i(X_s^\diamond) \Gamma_{jl}^r + (E_j^{\text{AI}} \zeta_r^i)_{X_s^\diamond} \right) E_r^{\text{AI}}(X_s^\diamond), \frac{\text{Log}_{X_s^\diamond}^{\text{AI}} V}{T-s} \right) ds
\end{aligned}$$

$$+ \sum_{i=1}^d \int_0^t \frac{\sum_{j=1}^d (\zeta_j^i(X_s^\diamond))^2}{2(T-s)} ds.$$

Substituting into equation (S14), we get

$$\begin{aligned} \frac{d\mathbb{P}_t}{d\mathbb{P}_t^\diamond}(X_{[0:t]}^\diamond) &= \exp \left[- (f(X_t^\diamond) - f(X_0^\diamond)) \right. \\ &+ \int_0^t \frac{ds}{2\sigma^2(T-s)} \left\{ \theta g_{X_s^\diamond}^{\text{AI}} \left(\text{Log}_{X_s^\diamond}^{\text{AI}} M, \text{Log}_{X_s^\diamond}^{\text{AI}} V \right) + \sigma^2 \sum_{i,j=1}^d (\zeta_j^i(X_s^\diamond))^2 \right. \\ &+ \left. \left. \sigma^2 \sum_{i,j,r=1}^d g_{X_s^\diamond}^{\text{AI}} \left(\zeta_j^i(X_s^\diamond) \left(\sum_{l=1}^d \zeta_l^i(X_s^\diamond) \Gamma_{jl}^r + (E_j^{\text{AI}} \zeta_r^i)_{X_s^\diamond} \right) E_r^{\text{AI}}(X_s^\diamond), \text{Log}_{X_s^\diamond}^{\text{AI}} V \right) \right\} \right] \\ &= \frac{\exp(f(X_0^\diamond))}{\exp(f(X_t^\diamond))} \exp \left\{ \Phi(t, X_{[0:t]}^\diamond) + \phi(t, X_{[0:t]}^\diamond) \right\}. \end{aligned}$$

Using the result from Schauer et al. [2017], that is

$$\frac{d\mathbb{P}_t^*}{d\mathbb{P}_t}(X_{[0:t]}^\diamond) = \frac{p(t, X_t^\diamond; T, V)}{p(0, U; T, V)},$$

we can easily get the desired result as equation (S10). \square

Since the functions ζ , the coefficients with respect to the basis $\mathfrak{B}_d^{\text{AI}}$ in the expression of the horizontal lift in local coordinates, cannot be expressed explicitly (as mentioned in Proposition S3), their approximation will be discussed below.

Corollary S1 (Approximation of functions ζ). *We approximate $\zeta_j^i(\gamma(t))$ when t is close to 0 under two special forms of $\gamma(t)$.*

1. Fix an integer $l \in \{1, \dots, d\}$, and consider the geodesic curve on $\mathcal{SP}(n)$ equipped with the AI metric such that $\gamma(0) = P$, $\dot{\gamma}(0) = P^{1/2} S_l P^{1/2}$, i.e. $\gamma(t) = P^{1/2} \exp(t S_l) P^{1/2}$. Suppose u_t is the unique horizontal lift defined in Proposition S3,

$$\Rightarrow (E_r^{\text{AI}} \zeta_j^i)_P = (d\zeta_j^i)_P (E_r^{\text{AI}}(P)) = \left. \frac{d}{dt} (\zeta_j^i \circ \gamma)(t) \right|_{t=0} \quad (1 \leq r \leq d).$$

With $\alpha_r(P) = \delta_{rl}$, Proposition S3 implies

$$\begin{aligned} (E_l^{\text{AI}} \zeta_j^i)_P &= - \sum_{k=1}^d \zeta_k^i(P) \Gamma_{lk}^j = - \sum_{k=1}^d \delta_{ik} \Gamma_{lk}^j = -\Gamma_{li}^j \quad (1 \leq i, j \leq d) \\ &\Rightarrow \left. \frac{d}{dt} (\zeta_j^i \circ \gamma)(t) \right|_{t=0} = -\Gamma_{li}^j \end{aligned}$$

Since $(\zeta_j^i \circ \gamma)(0) = \delta_{ij}$, $\lim_{t \rightarrow 0} \frac{(\zeta_j^i \circ \gamma)(t) - \delta_{ij}}{t} = -\Gamma_{li}^j$, then $\zeta_j^i(\gamma(t)) \approx \delta_{ij} - t \Gamma_{li}^j$ ($0 \leq t \ll 1$). Thus, for all $1 \leq i, j \leq d$: $\zeta_j^i(P^{1/2} \exp(t S_l) P^{1/2}) \approx \delta_{ij} - t \Gamma_{li}^j$ ($0 \leq t \ll 1$).

2. Consider an arbitrary $e = \sum_{l=1}^d \epsilon_l e_l \in \mathbb{R}^d$ such that $e_l \neq 0$ for $l \in \mathfrak{J} \subseteq \{1, \dots, d\}$ and let $\gamma(t) = \text{Exp}_P^{\text{AI}}(t u_0(e))$, then

$$\dot{\gamma}(0) = \sum_{l=1}^d \epsilon_l E_l^{\text{AI}}(P) = P^{1/2} \left(\sum_{l=1}^d \epsilon_l S_l \right) P^{1/2} \Rightarrow \alpha_r(X_0) = e_r \neq 0 \text{ for } r \in \mathfrak{J}.$$

$$\text{So, } \sum_{r \in \mathfrak{J}} \epsilon_r (E_r^{AI} \zeta_j^i)_{\gamma(t)} = - \sum_{r \in \mathfrak{J}} \epsilon_r \sum_{k=1}^d \zeta_k^i(\gamma(t)) \Gamma_{rk}^j.$$

$$\Rightarrow \sum_{r \in \mathfrak{J}} \epsilon_r (E_r^{AI} \zeta_j^i)_P = - \sum_{r \in \mathfrak{J}} \epsilon_r \sum_{k=1}^d \zeta_k^i(P) \Gamma_{rk}^j = - \sum_{r \in \mathfrak{J}} \epsilon_r \Gamma_{ri}^j.$$

On the other hand, setting the vector field $V = \sum_{l=1}^d \epsilon_l E_l^{AI} = \sum_{r \in \mathfrak{J}} \epsilon_r E_r^{AI} \in \Gamma(TSP(n))$, then we have

$$\begin{aligned} (V \zeta_j^i)_{\gamma(t)} &= \sum_{r \in \mathfrak{J}} \epsilon_r (E_r^{AI} \zeta_j^i)_{\gamma(t)} \text{ and } (V \zeta_j^i)_P = (d \zeta_j^i)_P(V_P) = \left. \frac{d}{dt} (\zeta_j^i \circ \gamma)(t) \right|_{t=0} \\ &\Rightarrow \lim_{t \rightarrow 0} \frac{(\zeta_j^i \circ \gamma)(t) - \delta_{ij}}{t} = \sum_{r \in \mathfrak{J}} \epsilon_r (E_r^{AI} \zeta_j^i)_P = - \sum_{r \in \mathfrak{J}} \epsilon_r \Gamma_{ri}^j. \end{aligned}$$

Thus, for all $1 \leq i, j \leq d$:

$$\zeta_j^i \left(P^{1/2} \exp\left(t \sum_{l=1}^d \epsilon_l S_l\right) P^{1/2} \right) \approx \delta_{ij} - t \sum_{l=1}^d \epsilon_l \Gamma_{li}^j \mathbf{I}_{l \in \mathfrak{J}},$$

for $0 \leq t \ll 1$, i.e. small t . Here \mathbf{I} stands for the indicator function.

Using Corollary S1, we are ready to approximate ϕ and Φ in Theorem S2.

Corollary S2. If X_t^\diamond is simulated using the piecewise approximation method in equation (S4) with sufficiently small time width, that is we have a simulation path $\{X_{t_k}^\diamond = y_{t_k}^\diamond\}_{k=0}^{m+1}$ such that $\max\{t_{k+1} - t_k\}_{k=0}^m$ is sufficiently small and $0 = t_0 < \dots < t_{m+1} = t$, then approximations of the functions ϕ in equation (S12) and Φ in equation (S13) are given by

$$\phi(t, X_{[0:t]}^\diamond) \approx \frac{d}{2} \log \frac{T-t}{T}, \quad (\text{S15})$$

$$\begin{aligned} \Phi(t, X_{[0:t]}^\diamond) &\approx \sum_{j=0}^m \frac{t_{k+1} - t_k}{T - t_k} \left[\frac{\theta \langle \log \{ (y_{t_k}^\diamond)^{-1/2} M (y_{t_k}^\diamond)^{-1/2} \}, \log \{ (y_{t_k}^\diamond)^{-1/2} V (y_{t_k}^\diamond)^{-1/2} \} \rangle_F}{\sigma^2} \right. \\ &\quad \left. + \frac{\langle \Gamma, \log((y_{t_k}^\diamond)^{-1/2} V (y_{t_k}^\diamond)^{-1/2}) \rangle_F}{2} \right], \quad (\text{S16}) \end{aligned}$$

with $\Gamma = \sum_{i,r=1}^d \Gamma_{ii}^r S_r$ and Christoffel symbols Γ_{ii}^r are given in Lemma S1.

Proof. Using the result of approximation for ζ_j^i in Corollary S1, we have for $0 \leq k \leq m$ and $1 \leq i, j \leq d$:

$$\zeta_j^i(X_{t_k}^\diamond) \approx \delta_{ij} \text{ and } U_t^\diamond(e_i) \approx E_i^{AI}(x_{t_k}^\diamond) \Rightarrow E_l^{AI} \zeta_j^i = 0 \text{ for all } 1 \leq l \leq d.$$

Thus,

$$\begin{aligned} \phi(t, X_{[0:t]}^\diamond) &= \sum_{i,j=1}^d \int_0^t \frac{(\zeta_j^i(X_s^\diamond))^2}{2(T-s)} ds \approx \sum_{i=1}^d \int_0^t \frac{1}{2(T-s)} ds = \frac{d}{2} \log \frac{T-t}{T}, \\ \Phi(t, X_{[0:t]}^\diamond) &= \int_0^t \left\{ \theta g_{X_s^\diamond}^{AI} \left(\text{Log}_{X_s^\diamond}^{AI} M, \frac{\text{Log}_{X_s^\diamond}^{AI} V}{\sigma^2(T-s)} \right) \right. \\ &\quad \left. + \sum_{i,j,r=1}^d g_{X_s^\diamond}^{AI} \left(\zeta_j^i(X_s^\diamond) \left(\sum_{l=1}^d \zeta_l^i(X_s^\diamond) \Gamma_{jl}^r + (E_j^{AI} \zeta_r^i)_{X_s^\diamond} \right) E_r^{AI}(X_s^\diamond), \frac{\text{Log}_{X_s^\diamond}^{AI} V}{2(T-s)} \right) \right\} ds \end{aligned}$$

$$\begin{aligned}
&\approx \int_0^t \frac{\theta \langle \log((X_s^\diamond)^{-1/2} M (X_s^\diamond)^{-1/2}), \log((X_s^\diamond)^{-1/2} V (X_s^\diamond)^{-1/2}) \rangle_F ds}{\sigma^2(T-s)} \\
&+ \sum_{i,j,r=1}^d \sum_{k=0}^m \frac{\langle \delta_{ij} \left(\sum_{l=1}^d \delta_{il} \Gamma_{jl}^r \right) S_r, \log(y_{t_k}^\diamond)^{-1/2} V (y_{t_k}^\diamond)^{-1/2} \rangle_F}{2(T-t_k)} (t_{k+1} - t_k) \\
&\approx \sum_{k=0}^m \frac{t_{k+1} - t_k}{T - t_k} \left[\frac{\theta \langle \log \{ (y_{t_k}^\diamond)^{-1/2} M (y_{t_k}^\diamond)^{-1/2} \}, \log \{ (y_{t_k}^\diamond)^{-1/2} V (y_{t_k}^\diamond)^{-1/2} \} \rangle_F}{\sigma^2} \right. \\
&\quad \left. + \frac{\langle \Gamma, \log((y_{t_k}^\diamond)^{-1/2} V (y_{t_k}^\diamond)^{-1/2}) \rangle_F}{2} \right]
\end{aligned}$$

□

3. STOCHASTIC COMPLETENESS ON $\mathcal{SP}(n)$

Before proving stochastic completeness, let us state the calculation of the Ricci curvature by Pennek [2020] in the case of the AI metric as well as the theorem in Hsu [2002] showing that manifolds having a lower bound on the Ricci curvature are stochastically complete.

Lemma S2. [Pennek, 2020] For any $P \in \mathcal{SP}(n)$, the Ricci curvature is given in terms of the basis $\mathfrak{B}_d^{AI}(P)$, as defined in equation (S7) :

$$Ric_P = -\frac{n}{4} \begin{pmatrix} I_n - \frac{1}{n} \mathbb{1}_{n,n} & 0 \\ 0 & I_{n(n-1)/2} \end{pmatrix},$$

where $\mathbb{1}_{n,n}$ is an $n \times n$ matrix, that has all entries equal to one.

Theorem S3. [Hsu, 2002] Consider a complete Riemannian manifold \mathcal{M} of dimension d , a fixed point $p \in \mathcal{M}$ and denote $d(x, p)$ as the distance between $x \in \mathcal{M}$ and p . Suppose that a negative, non-decreasing, continuous function $\kappa : [0, \infty) \rightarrow \mathbb{R}_{<0}$ satisfies

$$\kappa(r) \leq \frac{1}{d-1} \inf_{x \in \mathcal{M}} \{ Ric_{\mathcal{M}}(x) : d(x, p) = r \}$$

where $Ric_{\mathcal{M}}(x) = \{ Ric(X, X) : X \in T_x \mathcal{M} \text{ and } |X| = 1 \}$. If

$$\int_c^\infty \frac{1}{\sqrt{-\kappa(r)}} dr = \infty$$

for some constant c then \mathcal{M} is stochastically complete.

Proposition S4 (stochastic completeness). The Riemannian manifold $\mathcal{SP}(n)$ is stochastically complete if it is equipped with either

- (i) the LE metric,
- (ii) or the AI metric.

Proof. (i) Firstly, we show that $\mathcal{SP}(n)$ equipped with the LE metric has null sectional curvature everywhere, i.e. for all $1 \leq i, j \leq d$ and $P \in \mathcal{SP}(n)$ we have $\nabla_{E_i^{\text{LE}}(P)} E_j^{\text{LE}}(P) = 0$, where $E_i^{\text{LE}}, E_j^{\text{LE}} \in \mathfrak{B}_d^{\text{LE}}$, as defined in equation (S6).

We have the result that $\mathcal{SP}(n)$ endowed with the LE metric is isometric to the $\mathcal{S}(n)$ endowed with a Euclidean metric (Frobenius inner product) through the matrix logarithm function, that is $\log : \mathcal{SP}(n) \rightarrow \mathcal{S}(n)$ is a diffeomorphism and for all $P \in \mathcal{SP}(n)$ we have

$$g_P^{\text{LE}}(S_1, S_2) = \langle d_P \log S_1, d_P \log S_2 \rangle_F \quad S_1, S_2 \in T_P \mathcal{SP}(n).$$

So $\nabla^{\mathcal{SP}(n)}$ is the pull-back connection of $\nabla^{\mathcal{S}(n)}$ by the matrix logarithm function, and we have for any $P \in \mathcal{SP}(n)$ and $1 \leq i, j \leq d$:

$$\Rightarrow \nabla_{E_j^{\text{LE}}(P)}^{\mathcal{SP}(n)} E_j^{\text{LE}}(P) = \nabla_{\{(d_P \log)^{-1} S_i\}}^{\mathcal{SP}(n)} \{(d_P \log)^{-1} S_j\} = (d_P \log)^{-1} \nabla_{S_i}^{\mathcal{S}(n)} S_j = 0.$$

Thus, the Ricci curvature tensor also vanishes everywhere, and the required result is a direct consequence of Theorem S3.

(ii) We aim to show that the Ricci curvature satisfies the condition in Theorem S3, which implies the required result.

Let us fix a point $P \in \mathcal{SP}(n)$, and vary some point $Q \in \mathcal{SP}(n)$ such that $d_{\text{AI}}^2(P, Q) = r$ for some $r > 0$. Consider a tangent vector $v \in T_Q \mathcal{SP}(n)$ such that $|v| = 1$, that is $v = \sum_{i=1}^d \nu_i E_i^{\text{AI}}(Q) \in T_Q \mathcal{SP}(n)$ and $\sum_{i=1}^d \nu_i^2 = 1$.

By Lemma S2, denoting the (i, j) th entry of the Ricci curvature tensor at Q in matrix form as $\text{Ric}_Q^{(i,j)}$, we have

$$\begin{aligned} \text{Ric}(v, v) &= \sum_{i,j=1}^d \nu_i \text{Ric}_Q^{(i,j)} \nu_j = -\frac{n-1}{4} \sum_{i=j=1}^n \nu_i^2 - \frac{n}{4} \sum_{i=j=n+1}^d \nu_i^2 + \frac{1}{4} \sum_{i \neq j}^n \nu_i \nu_j \\ &= -\frac{n}{4} \sum_{i=1}^d \nu_i^2 + \frac{1}{4} \sum_{i,j=1}^n \nu_i \nu_j = -\frac{n}{4} + \frac{1}{4} \sum_{i,j=1}^n \nu_i \nu_j \geq -\frac{n}{4} - \frac{n^2}{4} = -\frac{d}{2} \end{aligned}$$

The last inequality holds due to the fact that $\sum_{i=1}^d \nu_i^2 = 1$. Therefore, using Theorem S3, we can set $\kappa(r) = \frac{-d}{2(d-1)}$, which is clearly a negative, non-decreasing, continuous function.

$$\Rightarrow \int_c^\infty \frac{1}{\sqrt{-\kappa(r)}} dr = \sqrt{\frac{2(d-1)}{d}} \int_c^\infty 1 dr = \infty \quad \text{for some constant } c.$$

Thus, $\mathcal{SP}(n)$ endowed with the AI metric is stochastically complete. □

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